

Bayesian Estimation in mixed effects models using Monolix and mlxR

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Outline

- 1 Introduction
- 2 Estimation of the population parameters using Monolix
- 3 Estimation of the individual parameters using Monolix
- 4 Mlxtran: a powerful language for complex hierarchical models

What is a model?

A joint probability distribution!

$p(y, \psi, \theta)$: joint probability distribution of y , ψ and θ .

- $y = (y_{ij}, 1 \leq i \leq N, 1 \leq j \leq n_i)$ are the observations,
- $\psi = (\psi_i, 1 \leq i \leq N)$ are the individual parameters,
- θ is the vector of population parameters.

What is a joint probability distribution?

A hierarchical model

$$p(y, \psi, \theta) = p(y|\psi, \theta)p(\psi|\theta)p(\theta)$$

- $p(y|\psi, \theta)$: conditional distribution of the observations,
- $p(\psi|\theta)$ population distribution of the individual parameters,
- $p(\theta)$ the distribution of θ can represent
 - the *inter population variability*,
 - the *uncertainty* on θ ,
 - the *prior distribution* of the population parameters.

Some tasks

$$p(y, \psi, \theta) = p(y|\psi, \theta)p(\psi|\theta)p(\theta)$$

1) Estimation of the population parameters (only y is given)

- *Maximum Likelihood approach:* maximize $p(y|\theta) = \int p(y, \psi|\theta) d\psi$
- *Bayesian approach:* compute/maximize

$$p(\theta|y) = \frac{p(y|\theta)p(\theta)}{p(y)}$$

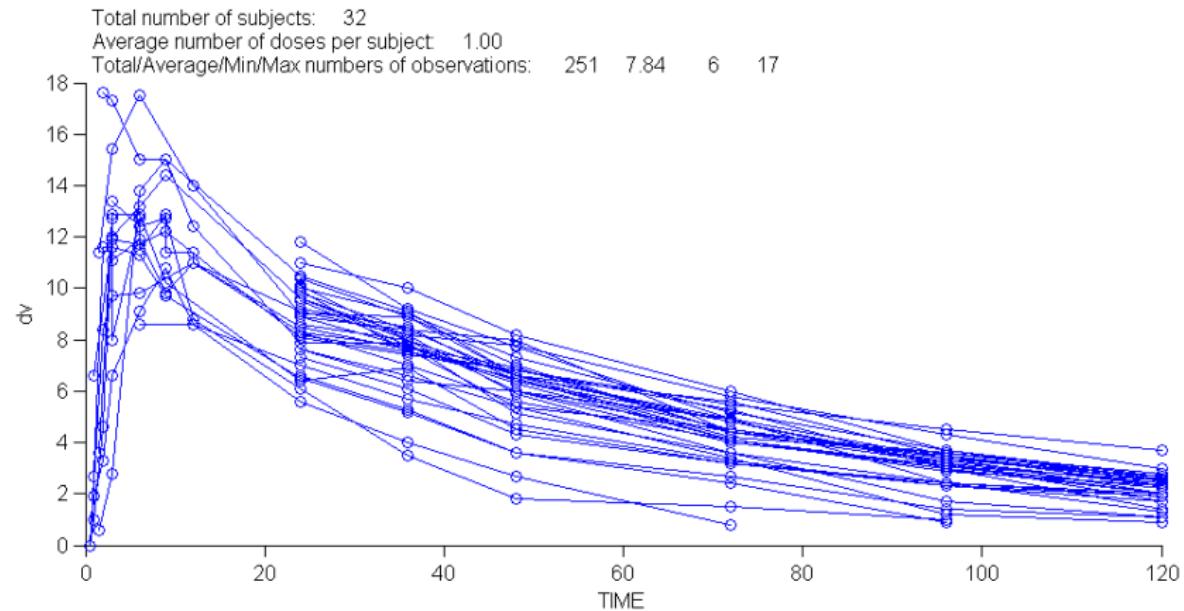
2) Estimation of the individual parameters (y and θ are given)

- *Individual approach:* maximize $p(y|\psi, \theta)$
- *Population approach:* compute/maximize

$$p(\psi|y, \theta) = \frac{p(y|\psi, \theta)p(\psi|\theta)}{p(y, \theta)}$$

A PK example

The warfarin PK data:



A PK example

- The PK model:

$$C(t, \psi) = \frac{D k_a}{V k_a - Cl} \left(e^{-(Cl/V)t} - e^{-k_a t} \right).$$

- The model for the observed concentrations:

$$y_{ij} = C(t_{ij}, \psi_i) + a \varepsilon_{ij}$$

i.e.

$$y_{ij} | \psi_i, a \sim \mathcal{N}(C(t_{ij}, \phi_i), a^2)$$

- The model for the individual PK parameters $\psi_i = (k_{a_i}, V_i, Cl_i)$:

$$\log(k_{a_i}) \sim \mathcal{N}(\log(k_{a_{\text{pop}}}), \omega_{k_a}^2)$$

$$\log(V_i) \sim \mathcal{N}(\log(V_{\text{pop}}), \omega_V^2)$$

$$\log(Cl_i) \sim \mathcal{N}(\log(Cl_{\text{pop}}), \omega_{Cl}^2).$$

II

Estimation of the population parameters using Monolix

Estimation of the population parameters

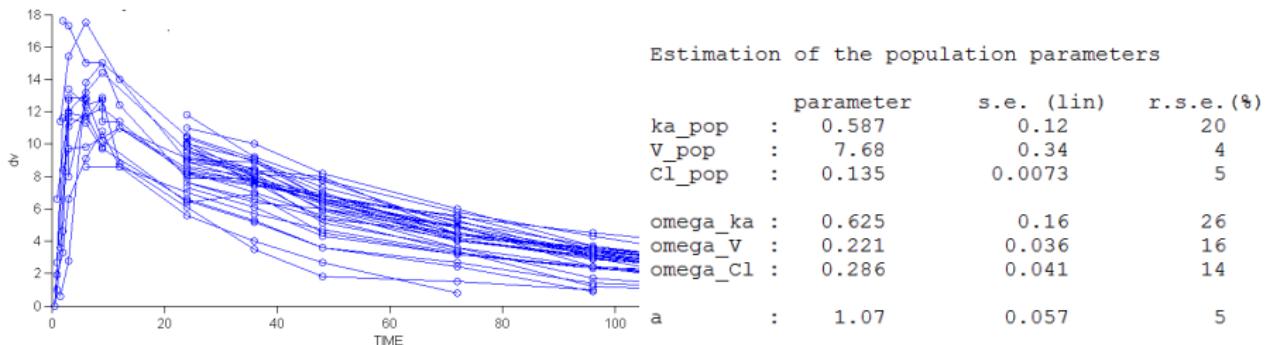
With Monolix, a given parameter can be

- **a fixed constant** if we have absolute confidence in its value or the data does not allow it to be estimated, essentially due to lack of identifiability.
- **estimated by maximum likelihood**, either because we have great confidence in the data or no information on the parameter.
- **estimated by introducing a prior** and calculating the maximum a posteriori (MAP) estimate or estimating the posterior distribution.

Estimation of the population parameters

ML estimation

The warfarin PK data provides a limited information about k_a



=> Let us introduce a prior for $k_{a\text{pop}}$:

$$\log(k_{a\text{pop}}) \sim \mathcal{N}(\log(k_{a^*}), \gamma^2)$$

Estimation of the population parameters

Monolix - 4.3.3 - warfarinPK_M0.mlxtran

Project Task Settings Workflow Graphics Test Tools ?

Run Assessment

The data and model

The data
warfarin_data.txt see

There is no covariate mixture

Distribution of the individual parameters L L L

Covariance model

1	0	0
0	1	0
0	0	1

Diagonal Full

The structural model
oral1_1cpt_kaVCI

Observation model

name	type	pred	error	r
y1	continuous	Cc	constant	

The initialization

Fixed effects
1 1 1

Stand. dev. of the random effects
1 1 1

Residual error parameters
1 0 1 0

Check initial fixed effects
Use the last estimates

The algorithms

New seed
123456

Numbers of iterations
K1 500
K2 200
 auto auto

Number of chains
2
 auto Min Size 50

Simulated Annealing

Monte-Carlo sizes
Pred. dist. 100
NPDE/VPC 500
LL 20000
Display 50

The results

Results folder
 Project name warfarinPK_M0
 User defined Browse

Standard errors
 Linearization
 Stoch. Approx.

Individual parameters
 Conditional modes
 Cond. means and s.d.

Log-likelihood
 Linearization
 Importance Sampling

Graphics
List

Estimation of the population parameters

Monolix - 4.3.3- warfarinPK_M0.mlxtran

Project Task Settings Workflow Graphics Test Tools ?

Run Assessment

The data and model

The data: warfarin_data.txt see mixture There is no covariate

Distribution of the individual parameters: L L L

Covariance model: 1 0 0
0 1 0
0 0 1 Diagonal Full

The structural model: oral1_1cpt_kaVCI

Observation model: name type pred error r
y1 continuous Cc constant

The initialization

Fixed effects: 1 Fixed Estimate Priors

Stand. dev. of the random effects: 1 1 1

Residual error parameters: 1 0 1 0

Check initial fixed effects
Use the last estimates

The algorithms

New seed: 123456
Numbers of iterations: K1 500 auto K2 200 auto
Number of chains: 2 auto Min Size 50
Simulated Annealing: checked

Monte-Carlo sizes: Pred. dist. 100 NPDE/VPC 500 LL 20000 Display 50

The results

Results folder: Project name warfarinPK_M0 User defined Browse

Standard errors: Linearization Stoch. Approx.

Individual parameters: Conditional modes Stoch. means and s.d.

Log-likelihood: Linearization Importance Sampling

Graphics: List

Estimation of the population parameters

Combining ML and MAP estimations

Monolix - 4.3.3- warfarinPK_M0.mlxtran

Project Task Settings Workflow Graphics Test Tools ?

Run Assessment

The data and model

warfarin_data.txt

There is no covariate

The data

Bayesian estimation

theta: intercept for ka

Method

M.A.P.

posterior distribution

Prior distribution

typical value (theta): 1

std. deviation (Z): 0.1

distribution: log-normal

theta= $\exp(Z)$

where "Z" is normally distributed

Accept Cancel

The structural model

oral1_1cpt_kaVCI

Observation model

name	type	pred	error	r
y1	continuous	Cc	constant	

The initialization

Fixed effects

1	1	1
---	---	---

Residual error parameters

1	0	1	0
---	---	---	---

The algorithms

New seed: 123456

Numbers of iterations: K1: 500, K2: 200, auto

Number of chains: 2, Min Size: 50, auto

Simulated Annealing: checked

Monte-Carlo sizes: Pred. dist.: 100, NPDE/VPC: 500, LL: 20000, Display: 50

The results

Results folder: warfarinPK_M0, User defined, Browse

Standard errors: Linearization, Stoch. Approx.

Individual parameters: Conditional modes, Cond. means and s.d.

Log-likelihood: Linearization, Importance Sampling

Graphics: List

Estimation of the population parameters

Combining ML and MAP estimations

Split θ into (θ_E, θ_M) where

- θ_E are the components of θ to be estimated with MLE,
- θ_M are the components of θ to be estimated with MAP.

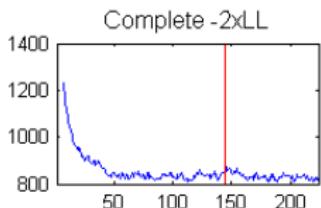
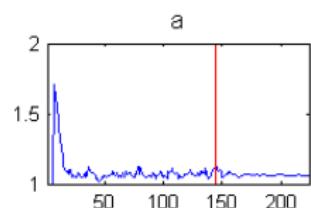
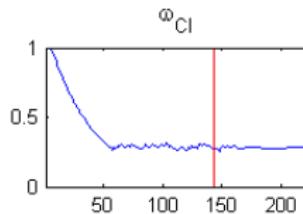
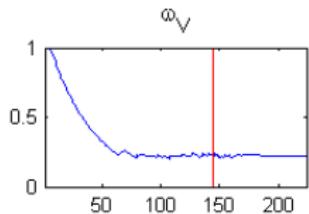
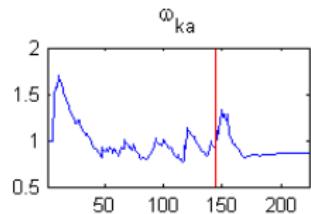
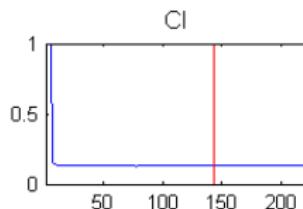
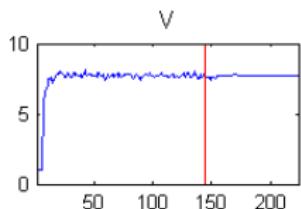
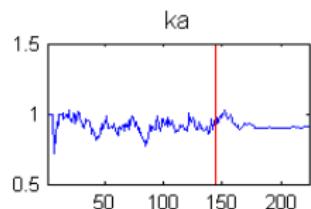
Then, $(\hat{\theta}_E, \hat{\theta}_M)$ maximizes the penalized likelihood of (θ_E, θ_M) :

$$\begin{aligned}(\hat{\theta}_E, \hat{\theta}_M) &= \arg \max_{\theta_E, \theta_M} \log(p(y, \theta_M; \theta_E)) \\&= \arg \max_{\theta_E, \theta_M} \{\mathcal{LL}_y(\theta_E, \theta_M) + \log(p(\theta_M))\},\end{aligned}$$

where $\mathcal{LL}_y(\theta_E, \theta_M) \stackrel{\text{def}}{=} \log(p(y|\theta_M; \theta_E)).$

Estimation of the population parameters

Combining ML and MAP estimations



Estimation of the population parameters

Combining ML and MAP estimations

The MAP estimate of $\theta = (ka_{\text{pop}}, V_{\text{pop}}, \dots)$ is the solution of the following minimization problem:

$$\hat{\theta}^{\text{MAP}} = \arg \min_{\theta} \left\{ -2\mathcal{LL}_y(\theta) + \frac{1}{\gamma^2} (\log(ka_{\text{pop}}) - \log(ka^*))^2 \right\}.$$

We can compute the MAP estimate of ka_{pop} for various values of γ and $ka^* = 1$:

γ	0	0.02	0.03	0.2	0.3	$+\infty$
$\hat{ka}_{\text{pop}}^{\text{MAP}}$	1	0.991	0.963	0.809	0.623	0.587

Estimation of the population parameters

Combining ML and posterior distribution estimation

Monolix - 4.3.3- warfarinPK_M3.mixtran

Project Task Settings Workflow Graphics Test Tools ?

Run Assessment

The data and model

The data warfarin_data.txt see mixture

There is no covariate

Distribution of the individual parameters L L L

Covariance model 1 1 0

The structural model oral1_fcpt_kaVCI

Observation model

name	type	pred	error	r
y1	continuous	Cc	constant	

The initialization

Fixed effects 1 1 1

Residual error parameters 1 0 1 0

Bayesian estimation

theta: intercept for ka

Method

M.A.P.

posterior distribution

Prior distribution

typical value (theta): 1

std. deviation (Z): 0.1

distribution: log-normal

theta= $\exp(Z)$
where "Z" is normally distributed

Accept Cancel

The algorithms

New seed 123456

Numbers of iterations K1 300 K2 200

Annealing Min Size 2 auto 50

Monte-Carlo sizes Pred. dist. 100 NPDEVPC 500 LL 20000 Display 50

The results

Results folder warfarinPK_M3

Standard errors

Linearization

Stoch. Approx.

Individual parameters

Conditional modes

Cond. means and s.d.

Log-likelihood

Linearization

Importance Sampling

Graphics List

Project name warfarinPK_M3

User defined Browse

Estimation of the population parameters

Combining ML and posterior distribution estimation

Split θ into (θ_E, θ_R) where

- θ_E are the components of θ to be estimated with MLE,
- θ_R are the components of θ whose posterior distribution is estimated.

Then,

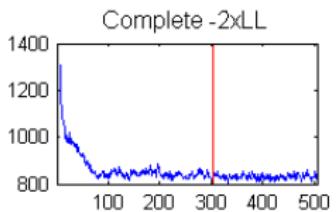
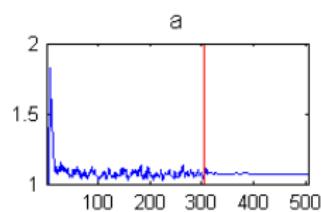
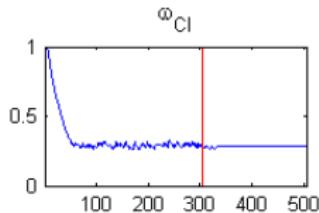
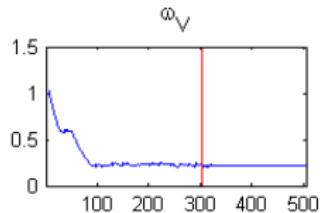
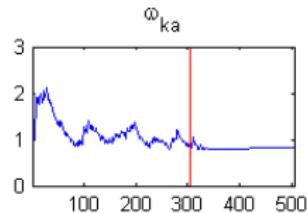
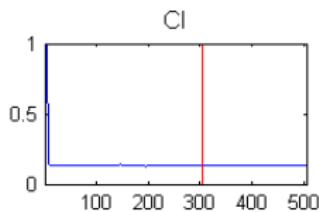
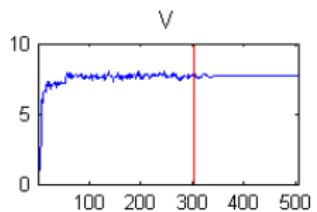
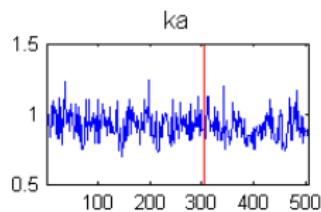
- Compute the maximum likelihood of θ_E :

$$\begin{aligned}\hat{\theta}_E &= \arg \max_{\theta_E} p(y; \theta_E) \\ &= \arg \max_{\theta_E} \int p(y, \theta_R; \theta_E) d\theta_R.\end{aligned}$$

- Estimate the conditional distribution $p(\theta_R|y; \hat{\theta}_E)$.

Estimation of the population parameters

Combining ML and posterior distribution estimation



Estimation of the population parameters

Combining ML and posterior distribution estimation

Monolix - 4.3.3- warfarinPK_M3.mixtran

Project Task Settings Workflow Graphics Test Tools ?

Run Assessment

List of graphics

Selection

All Reduced Simulation Other

Reduced

Project Summary	1
Spaghetti Plot	1
Individual Fits	1
Obs. vs Pred.	1
Covariates	1
Parameters Distribution	1
Random Effects (boxplot)	1
Random Effects (joint dist.)	1
Convergence SAEM	1

Simulation

Residuals	1
VPC	1
Time to Event data	1
NPC	1
BLQ	1
Prediction Distribution	1

Other

Categorized Data	1
Transition Probabilities	1
Posterior distribution	1
Contribution to likelihood	1

Save OK Cancel

The structural model
oral1_1cpt_kaVCI

parameters

Observation model

name	type	pred	error	r
y1	continuous	Cc	constant	<input type="checkbox"/>

Random effects

Residual error parameters

1	0	1	0
---	---	---	---

Monte-Carlo sizes

Pred. dist.	NPDE/VPC	LL	Display
100	500	20000	50

Results folder

Project name warfarinPK_M3

User defined Browse

Standard errors

Linearization

Stoch. Approx.

Individual parameters

Conditional modes

Cond. means and s.d.

Log-likelihood

Linearization

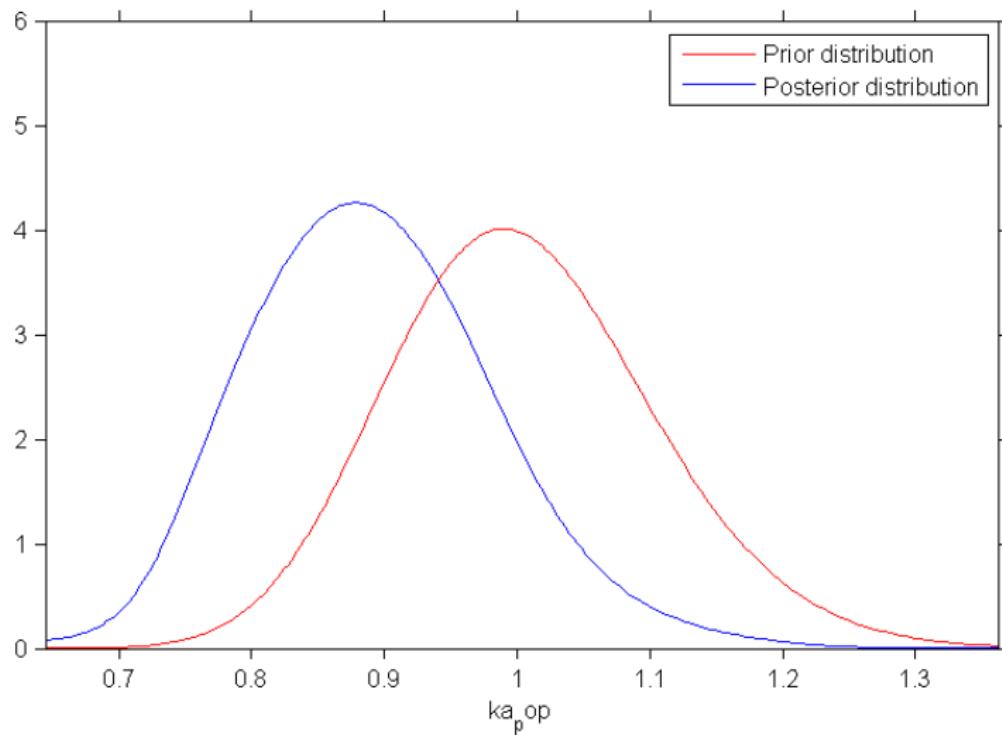
Importance Sampling

Graphics

List

Estimation of the population parameters

Combining ML and posterior distribution estimation



III

Estimation of the individual parameters using Monolix

Estimation of the individual parameters

Once θ has been estimated, Monolix allows one:

- to compute the maximum a posteriori (MAP) estimate of ψ_i , for $i = 1, 2, \dots, N$

$$\hat{\psi}_i^{\text{mode}} = \arg \max_{\psi_i} p(\psi_i | y_i; \hat{\theta})$$

- to estimate the conditional distribution $p(\psi_i | y_i; \hat{\theta})$, for $i = 1, 2, \dots, N$ using a Metropolis Hastings (MH) algorithm and estimate the conditional mean

$$\begin{aligned}\hat{\psi}_i^{\text{mean}} &= \widehat{\mathbb{E}(\psi_i | y_i; \hat{\theta})} \\ &= \frac{1}{K} \sum_{k=1}^K \psi_i^{(k)}.\end{aligned}$$

Estimation of the individual parameters

Conditional mode and conditional distribution

Monolix - 4.3.3- warfarinPK_M0 mlxtran

Project Task Settings Workflow Graphics Test Tools ?

Run Assessment

The data and model

The data: warfarin_data.txt see

The covariate model: mixture transform

wt: 0 0 0

Distribution of the individual parameters: L L L

Covariance model: 1 0 0
0 1 0
0 0 1
Diagonal Full

The structural model: oral1_1cpt_kaVCI

Observation model: name: y1 type: continuous pred: Cc error: constant r: 0

The initialization

Fixed effects: 1 1 1

Stand. dev. of the random effects: 1 1 1

Residual error parameters: 1 0 1 0

Check initial fixed effects
Use the last estimates

The algorithms

New seed: 123456

Numbers of iterations: K1: 500, K2: 200, auto

Number of chains: 2, Min Size: 50, auto

Simulated Annealing: checked

Monte-Carlo sizes: Pred. dist.: 100, NPDE/VPC: 500, LL: 20000

Display: 50

The results

Results folder: warfarinPK_M0
Project name: checked
User defined: uncheckable

Browse

Standard errors: Linearization checked, Stoch. Approx. uncheckable

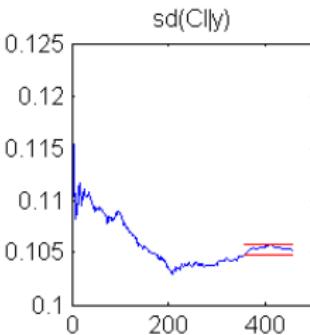
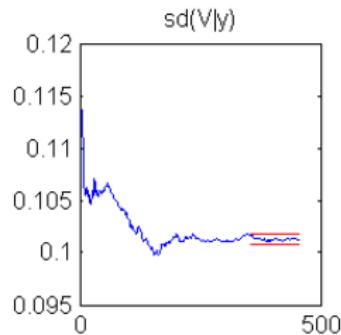
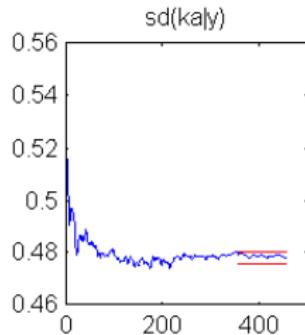
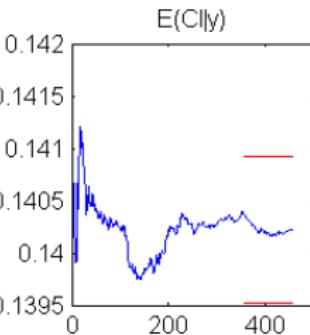
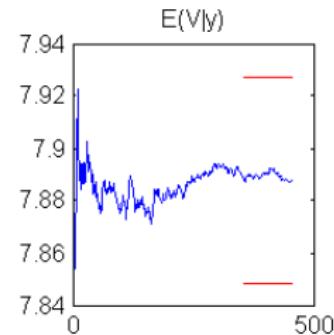
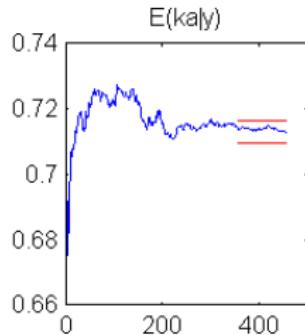
Individual parameters: Conditional modes checked, Cond. means and s.d. checked

Log-likelihood: Linearization checked, Importance Sampling uncheckable

Graphics: List

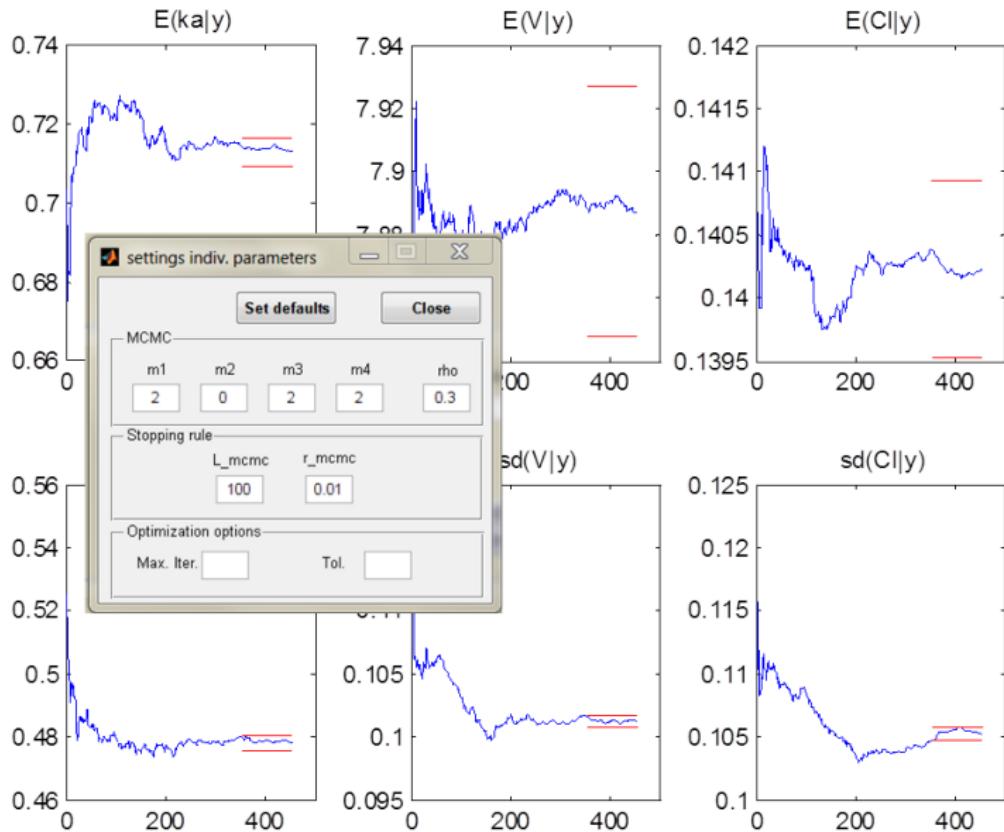
The Metropolis Hastings algorithm

Transitions kernels and stopping rule



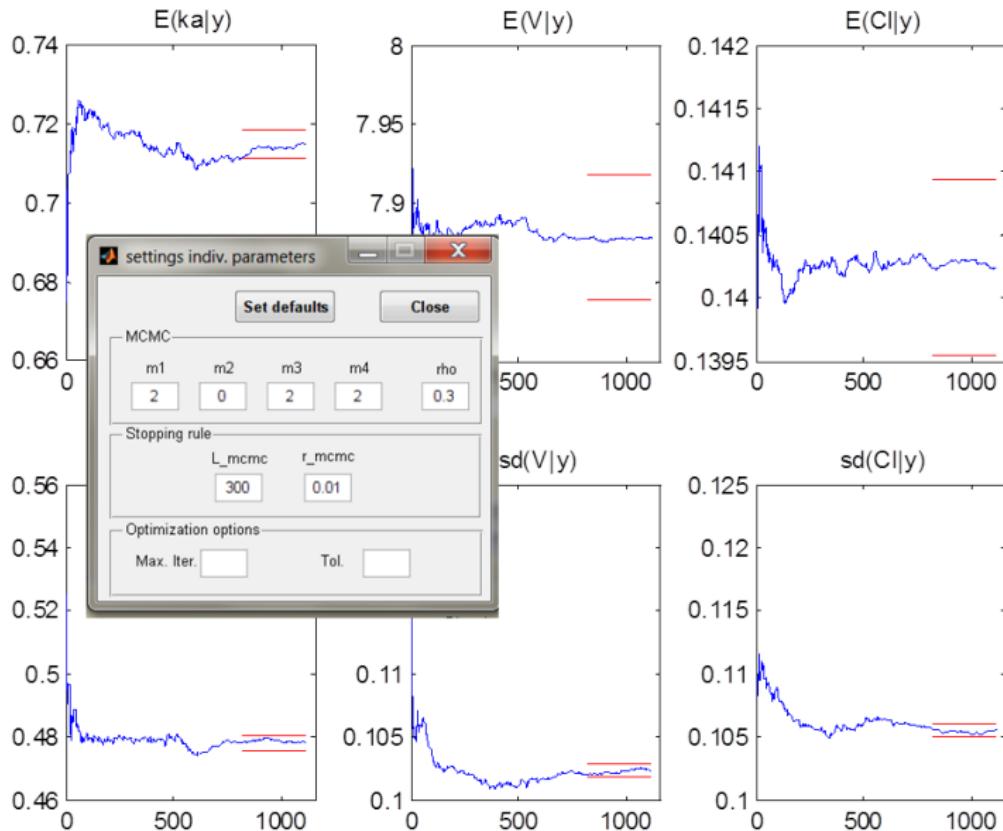
The Metropolis Hastings algorithm

Transitions kernels and stopping rule



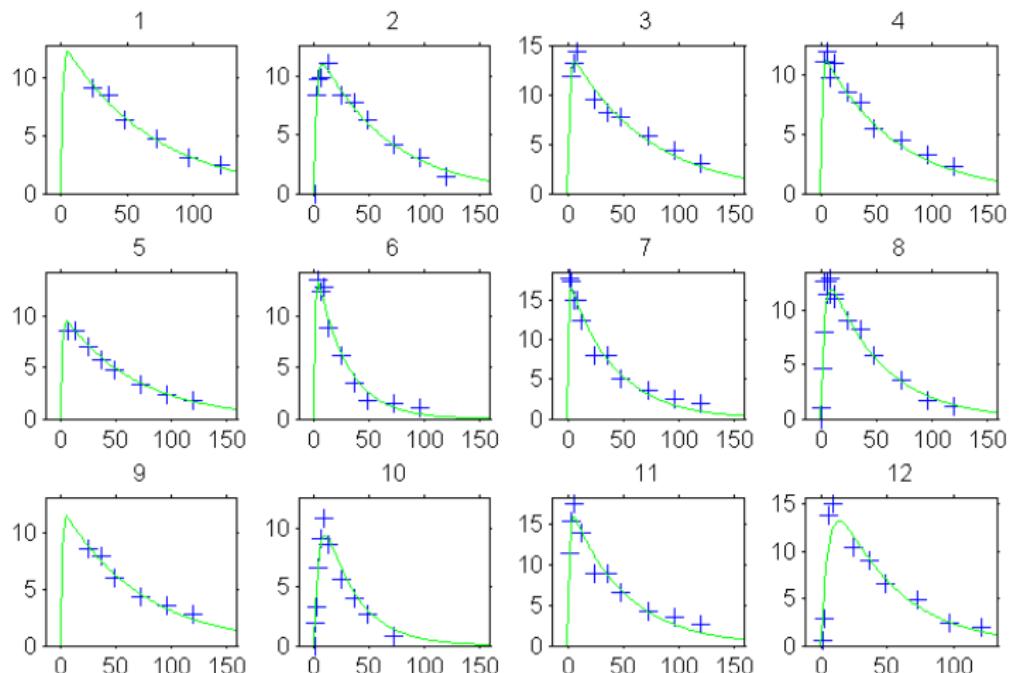
The Metropolis Hastings algorithm

Transitions kernels and stopping rule



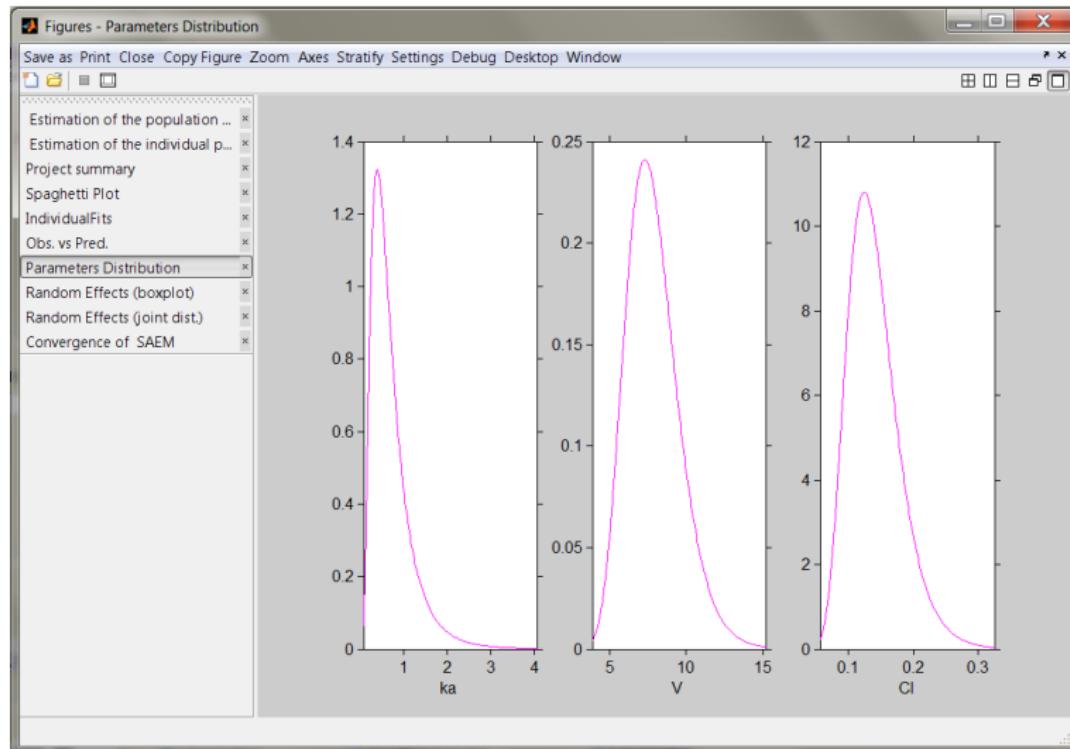
Some diagnostic plots

Individual fits



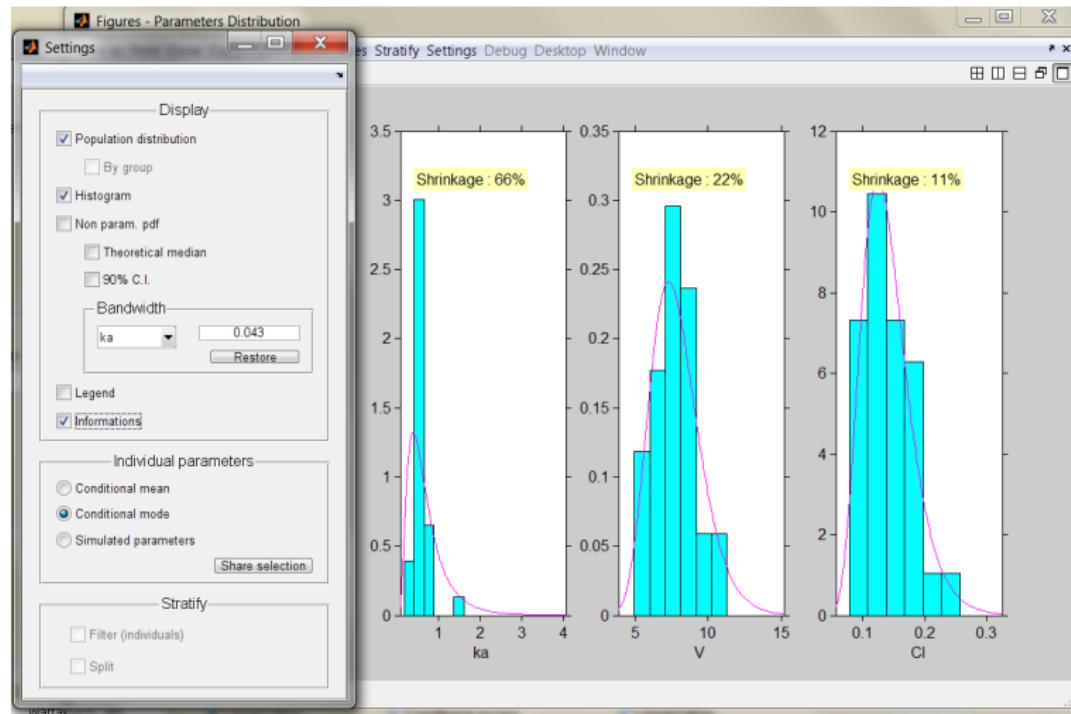
Some diagnostic plots

Distribution of the individual parameters



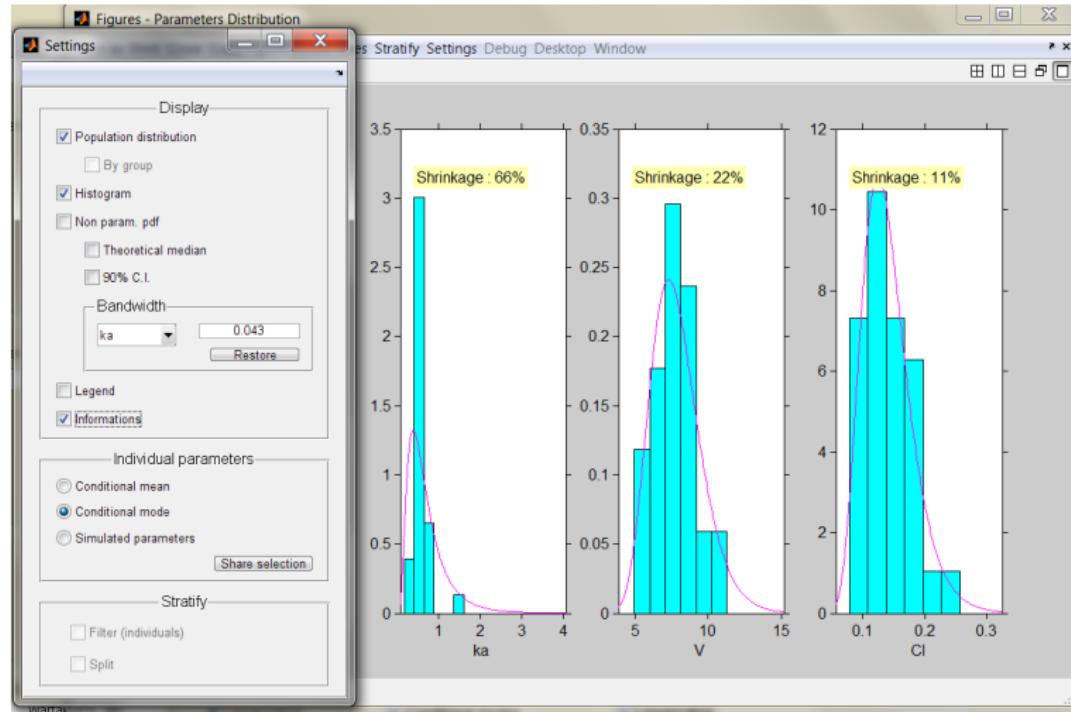
Some diagnostic plots

Distribution of the individual parameters



Some diagnostic plots

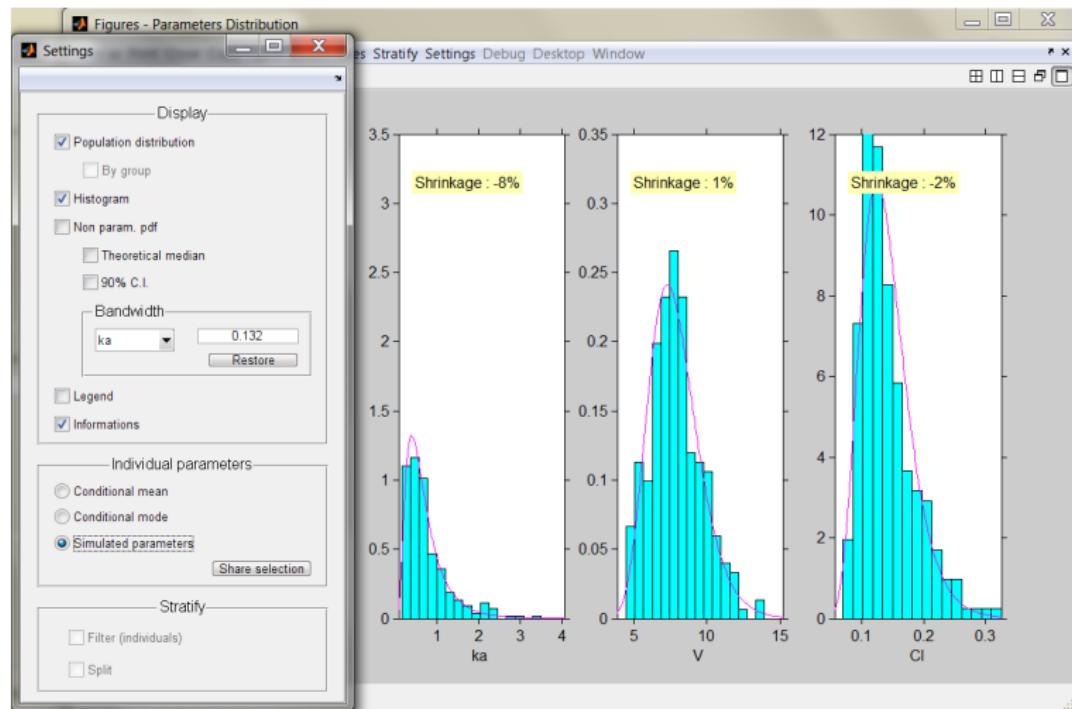
Distribution of the individual parameters



=> Don't use the MAP/conditional mode/posthoc/EBEs for diagnostic!

Some diagnostic plots

Distribution of the individual parameters



=> Use individual parameters simulated with the conditional distribution $p(\psi_i|y_i, \hat{\theta})$!

Some diagnostic plots

Using simulated individual parameters

Indeed, whenever data becomes sparse or uninformative, the EBE distribution will shrink towards the population value.

A particularly effective solution is to draw individual parameters $(\psi_i^{(k)}, 1 \leq k \leq K)$ with the conditional distribution $p(\psi_i|y_i, \hat{\theta})$ rather than taking the mode.

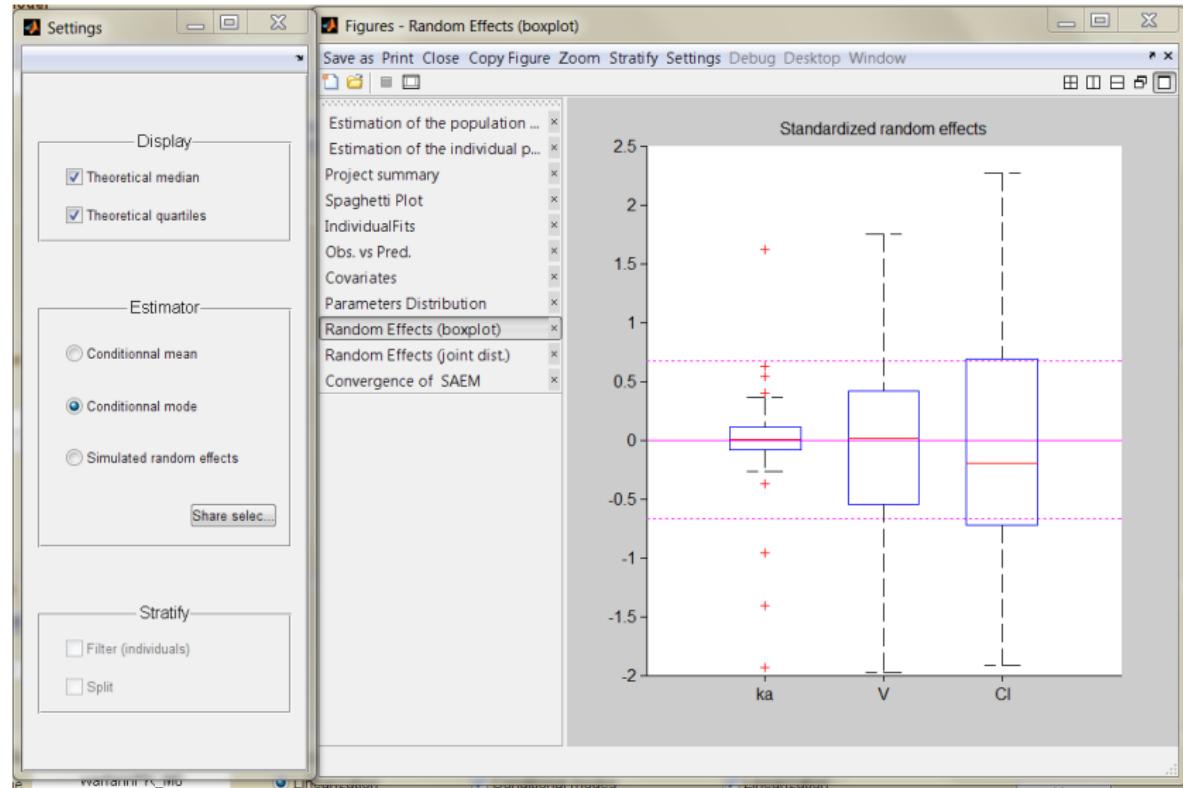
The resulting estimator is unbiased under H_0 in the following sense:

$$\begin{aligned} p(\psi_i) &= \int p(\psi_i|y_i)p(y_i)dy_i \\ &= \mathbb{E}_{y_i}(p(\psi_i|y_i)). \end{aligned}$$

This relationship is a fundamental one when considering inverse problems, incomplete data models, mixed effects models, etc.

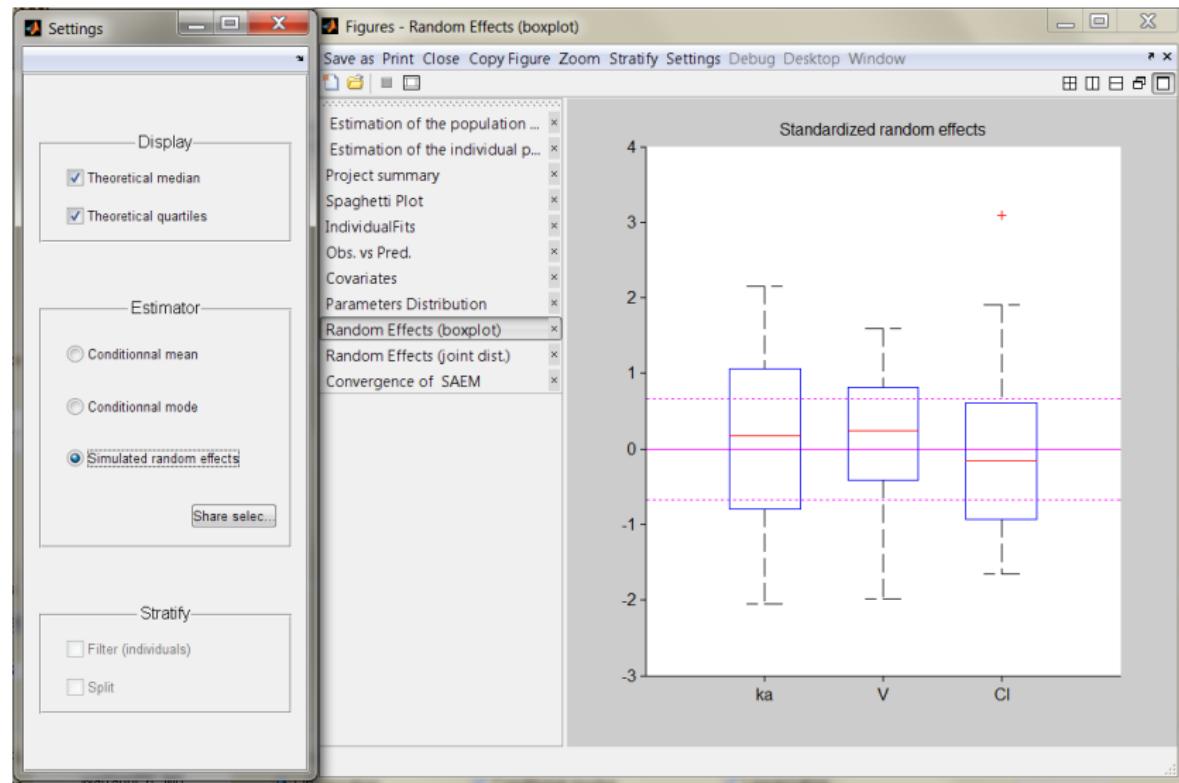
Diagnostic plots

Marginal distributions of the random effects



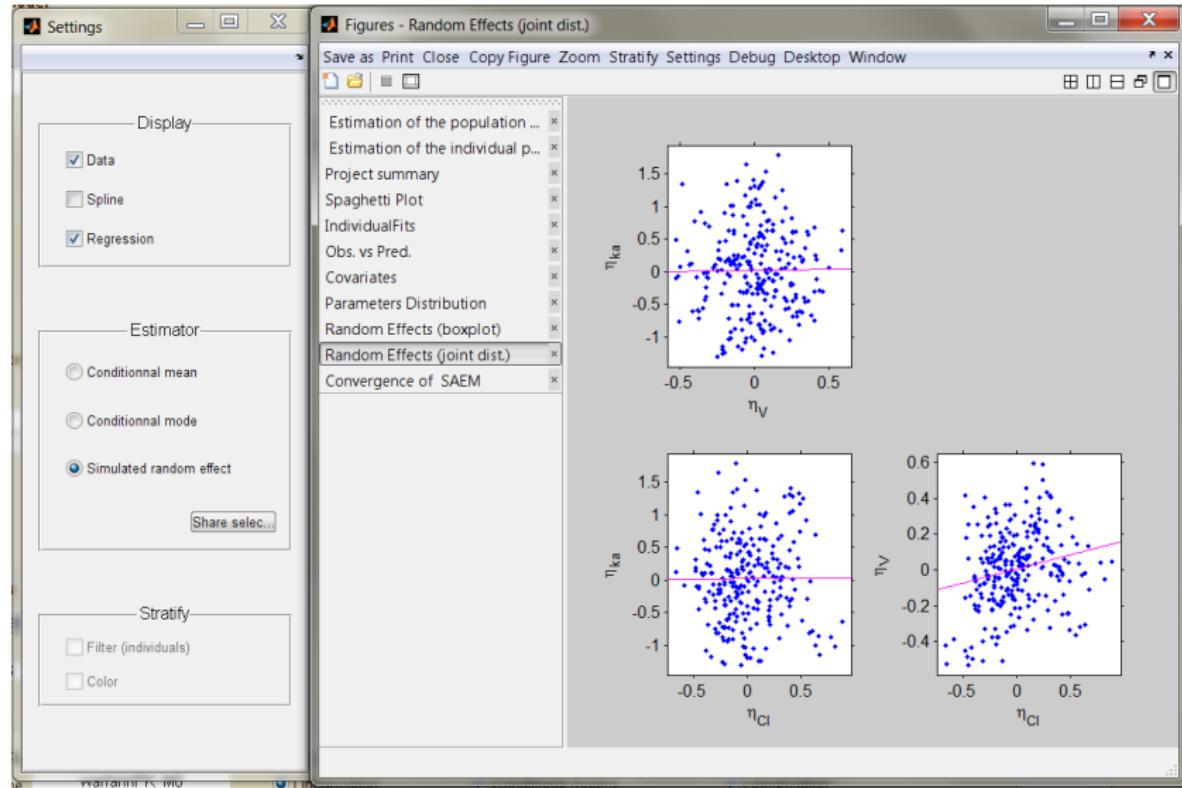
Diagnostic plots

Marginal distributions of the random effects



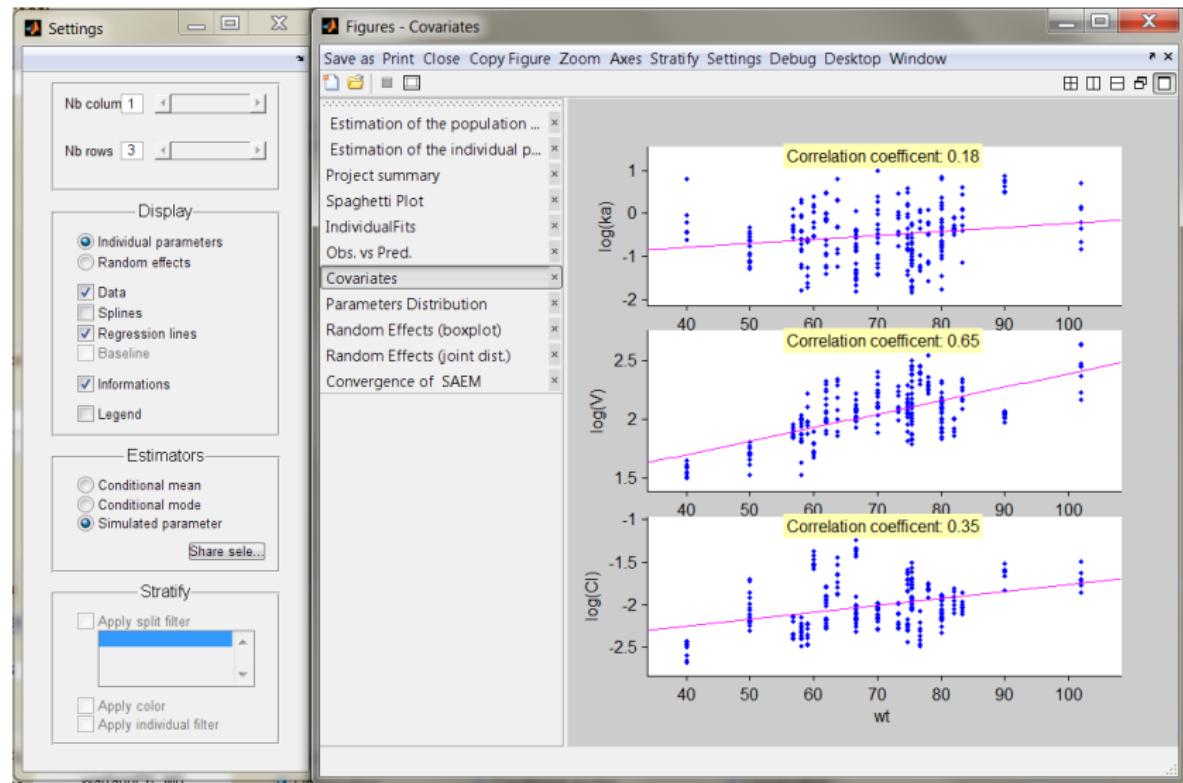
Diagnostic plots

Joint distributions of the random effects



Diagnostic plots

Relationship between parameters and covariates



IV

Mlxtran: a powerful
language for complex
hierarchical models

A hierarchical model

- The PK model:

$$C(t, \psi) = \frac{D k_a}{V k_a - Cl} \left(e^{-(Cl/V)t} - e^{-k_a t} \right).$$

- The model for the observed concentrations:

$$y_{ij} = C(t_{ij}, \psi_i) + a \varepsilon_{ij}$$

i.e.

$$y_{ij} | \psi_i, a \sim \mathcal{N}(C(t_{ij}, \phi_i), a^2)$$

- The model for the individual PK parameters $\psi_i = (k_{a_i}, V_i, Cl_i)$:

$$\log(k_{a_i}) \sim \mathcal{N}(\log(k_{a_{\text{pop}}}), \omega_{k_a}^2)$$

$$\log(V_i) \sim \mathcal{N}(\log(V_{\text{pop}}), \omega_V^2)$$

$$\log(Cl_i) \sim \mathcal{N}(\log(Cl_{\text{pop}}), \omega_{Cl}^2).$$

- The distribution for $k_{a_{\text{pop}}}$: $\log(k_{a_{\text{pop}}}) \sim \mathcal{N}(\log(k_a^*), \gamma^2)$

Mlxtran code for hierarchical model

```
[LONGITUDINAL]
input = {ka, V, Cl, a}
EQUATION:
C = pkmodel(ka, V, Cl)
DEFINITION:
y = {distribution = normal, prediction = C, sd = a}
;-----

[INDIVIDUAL]
input = {ka_pop, omega_ka, V_pop, omega_V, Cl_pop, omega_Cl}
DEFINITION:
ka = {distribution = lognormal, mean = log(ka_pop), sd = omega_ka}
V = {distribution = lognormal, mean = log(V_pop) , sd = omega_V }
Cl = {distribution = lognormal, mean = log(Cl_pop), sd = omega_Cl}
;-----

[POPULATION]
input = {ka_star, gamma_ka}
DEFINITION:
ka_pop = {distribution = lognormal, mean = log(ka_star), sd = gamma_ka}
```

R code using Simulx

```
p <- c(ka_star=0.5, gamma_ka=0.1, omega_ka=0.2,
      V_pop=10, omega_V=0.2,
      Cl_pop=1, omega_Cl=0.15,
      a=0.5)

f   <- list(name='C', time=seq(0, 30, by=0.1))
obs <- list(name='y', time=seq(1, 30, by=3))
ind <- list(name=c('ka', 'V', 'Cl'))
pop <- list(name='ka_pop')

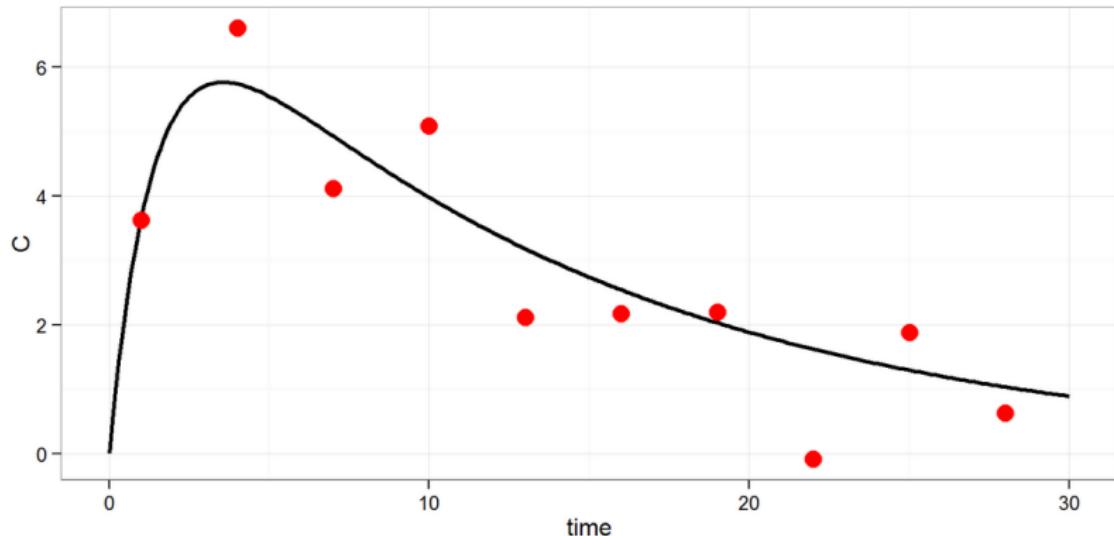
res <- simulx(model      = 'mlxtran1.txt',
               parameter = p,
               treatment = list(time=0, amount=100),
               output    = list(pop, ind, f, obs))

print(res$parameter)
```

```
##      ka_pop        ka         V         Cl
## 1 0.6039348 0.7340071 10.46294 1.327486
```

R code using Simulx

```
print(ggplotmlx() + geom_line(data=res$C, aes(x=time, y=C), size=1) +  
      geom_point(data=res$y, aes(x=time, y=y), colour='red', size=4))
```



R code using Simulx

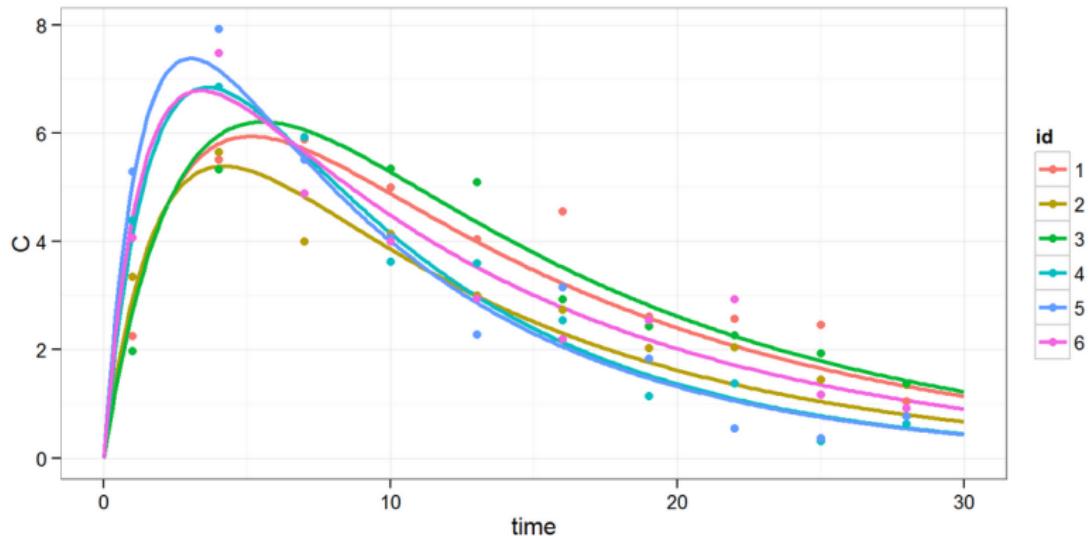
```
res <- simulx(model      = 'mlxtran1.txt',
               parameter = p,
               treatment = list(time=0, amount=100),
               output    = list(pop, ind, f, obs),
               group     = list(size=c(2,3), level=c('population','individual')))

print(res$parameter)
```

```
##   id   ka_pop       ka        V        Cl
## 1  1  0.4505751  0.4021775 11.465519  0.8554521
## 2  2  0.4505751  0.5028199 12.809041  1.1290427
## 3  3  0.4505751  0.3519262 10.550282  0.8050163
## 4  4  0.5421385  0.5509877  9.718898  1.0911209
## 5  5  0.5421385  0.7338368  9.671072  1.0744846
## 6  6  0.5421385  0.7349852 11.236472  0.8981349
```

R code using Simulx

```
print(ggplotmlx() +  
  geom_line( data=res$C, aes(x=time, y=C, colour=id), size=1) +  
  geom_point(data=res$y, aes(x=time, y=y, colour=id), size=2))
```



Simulx

A R function of the mlxR package
for computing predictions and sampling longitudinal data
from Mlxtran and PharmML models.

User Guide Videos Case studies Simulx & Shiny mlxR Notes Installation



User Guide

Learn how to use `simulx` with many illustrative examples



Case Studies

Discover several examples of practical use of `simulx`, including clinical trial simulation, modelling and simulation workflow...



Shiny

See how to combine `simulx` with Shiny and produce web applications and training material