

Surv-CRM-12

A Bayesian Phase I/II Survival CRM for Right-Censored Toxicity Endpoints with Competing Disease Progression

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October 26th, 2023

CHALLENGES FOR DOSE-FINDING TRIALS IN ONCOLOGY

- Delayed effects, including toxicities (**extended observation windows**)
 - Depending on the patient accrual rate, generate **incomplete data**

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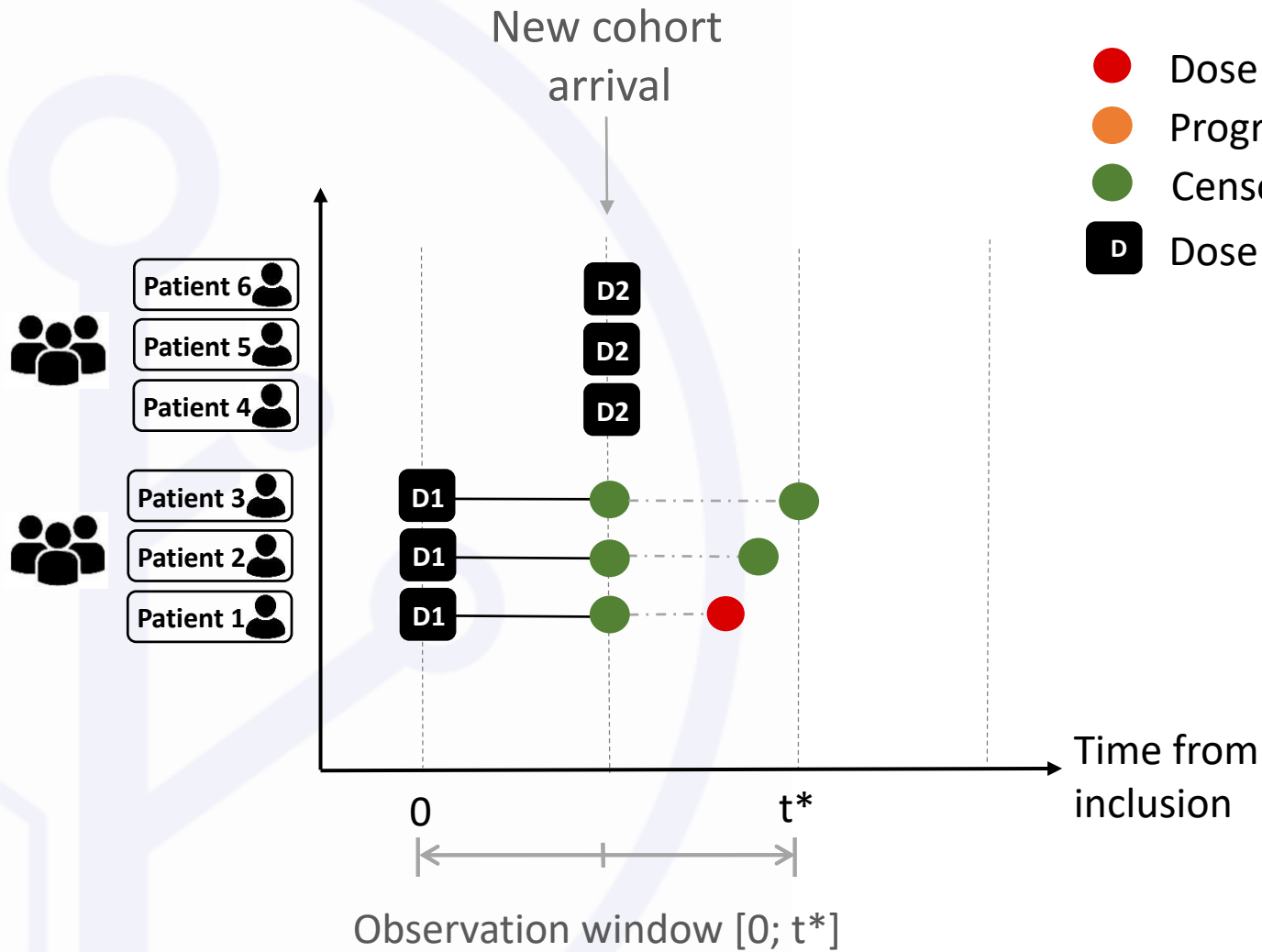
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- **Intercurrent events**: death, disease progression, consent withdrawal or physician discretion
 - **Alternative treatment** off the protocol will be proposed
 - Trial discontinuation, **precluding complete toxicity assessment**
 - Define a competing risks framework

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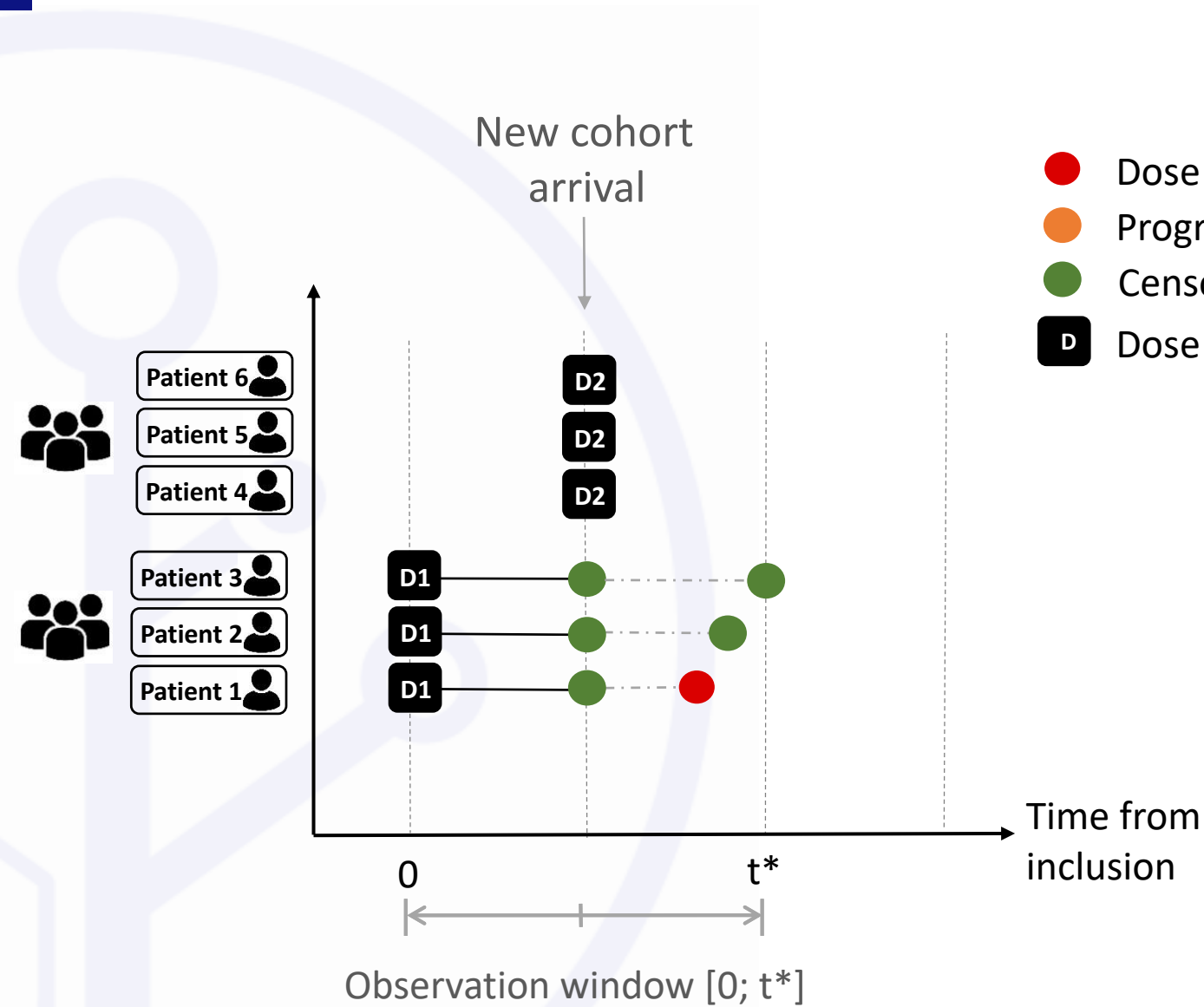
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 - Define a competing risks framework
- **New definition of the target dose** (instead of the MTD)

INCOMPLETE DATA

- Dose Limiting Toxicity (DLT)
- Progression
- Censored
- D Dose level



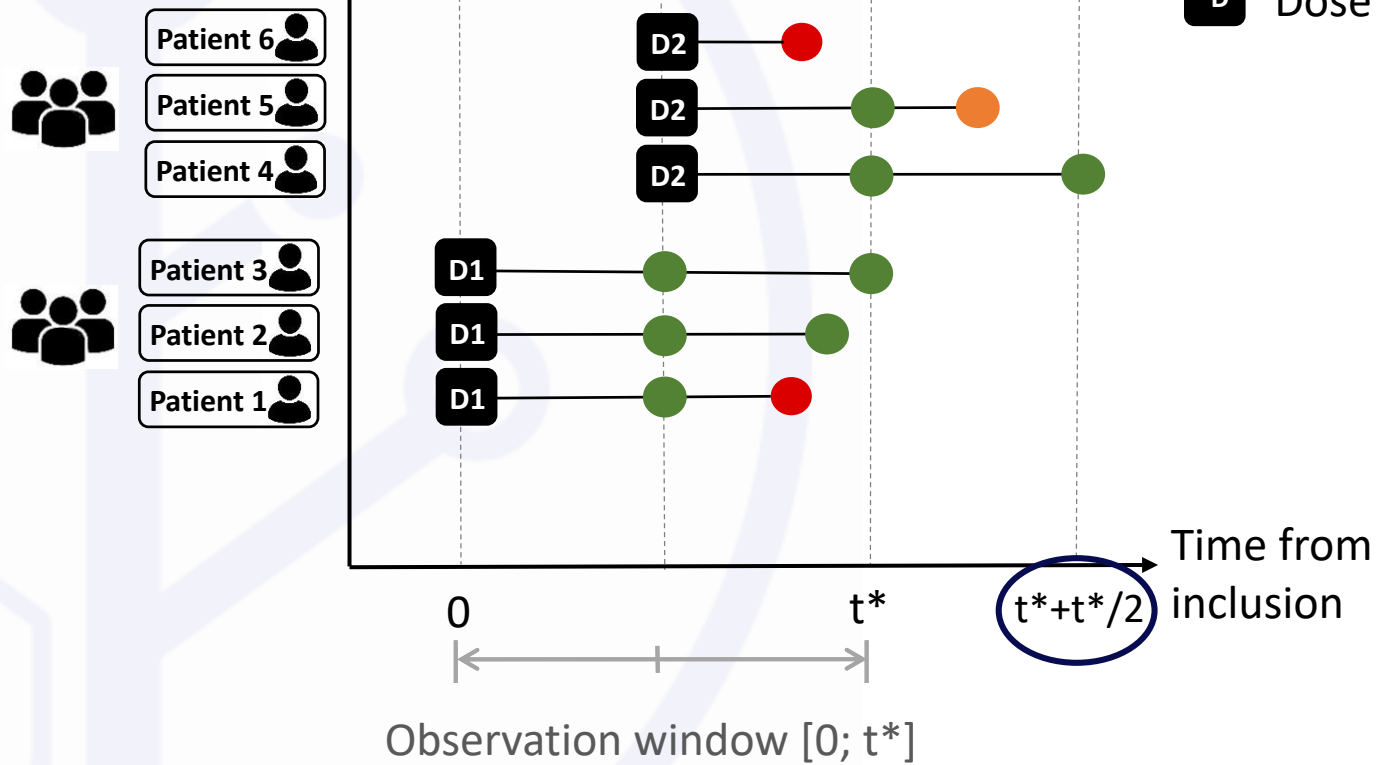
INCOMPLETE DATA



→ **Administrative censoring**
incomplete observation at the
time of dose-assignment

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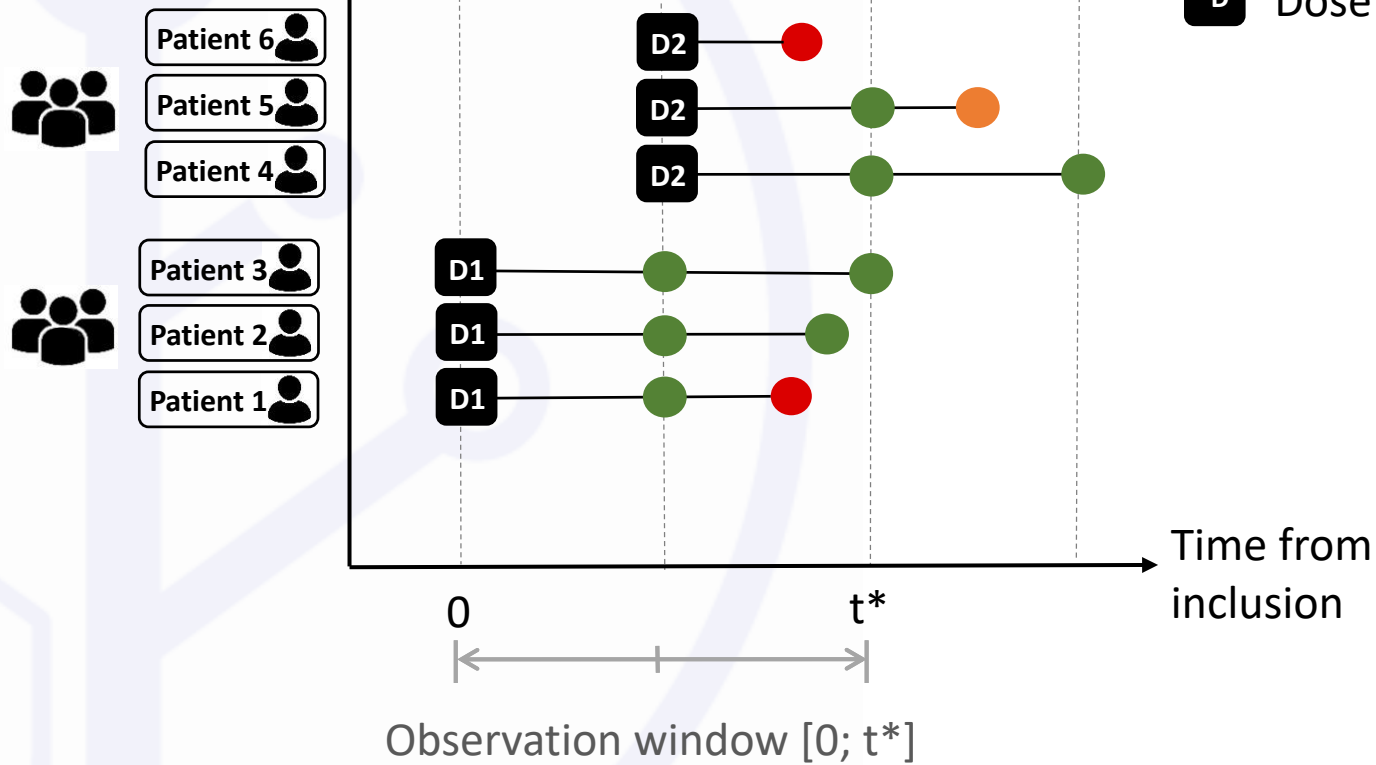
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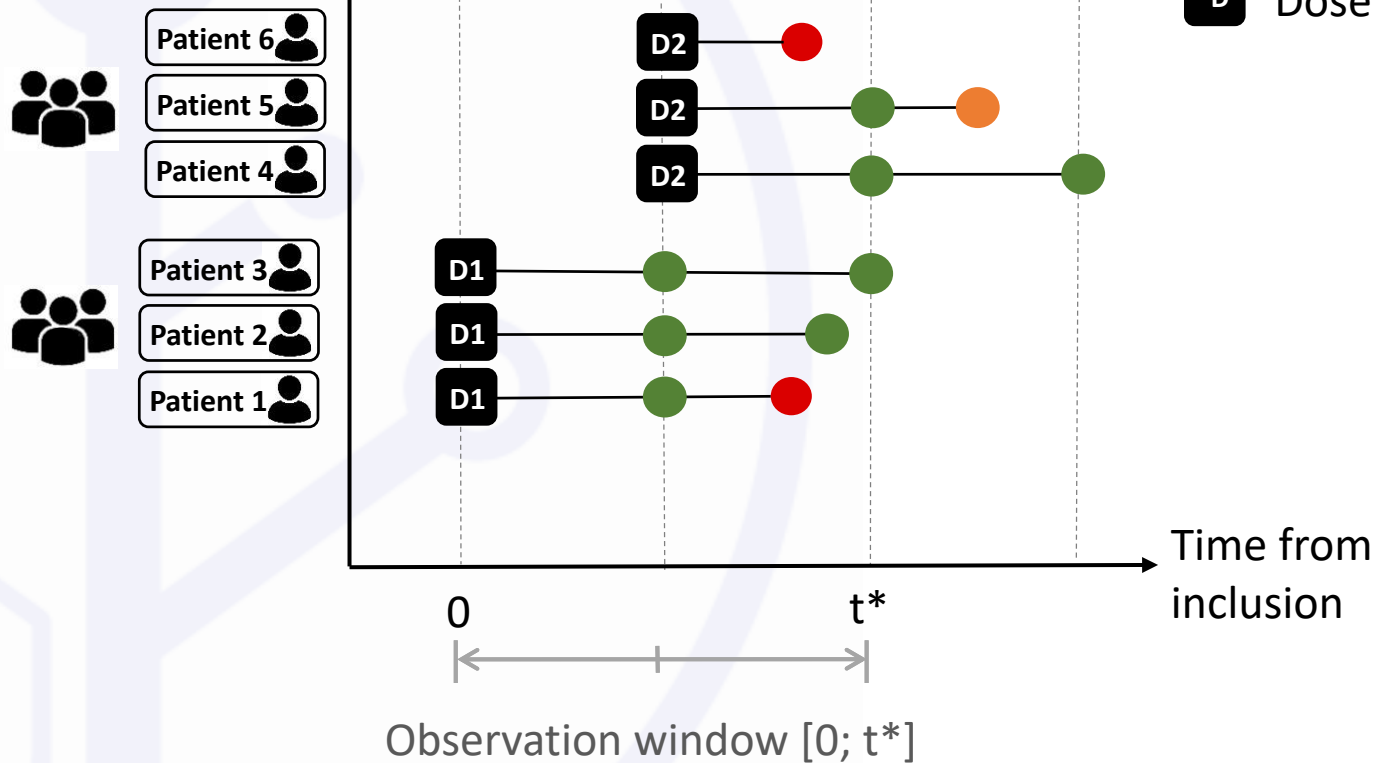


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→ **Informative censoring**
due to trial discontinuations

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CONTINUAL REASSESSMENT METHOD (CRM)
→ COMPLETE DATA

HANDLE ADMINISTRATIVE CENSORING IN PHASE I DESIGNS

- Mostly, weighting the observation with its partial follow-up time (**binary endpoints**)
 - TITE-CRM (Cheung and Chappell 2000), TITE-BOIN (Yuan et al. 2018), TITE-mTPI (Lin and Yuan 2020)

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- Mostly, weighting the observation with its partial follow-up time (**binary endpoints**)
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HANDLE INFORMATIVE CENSORING DUE TO TRIAL DISCONTINUATIONS

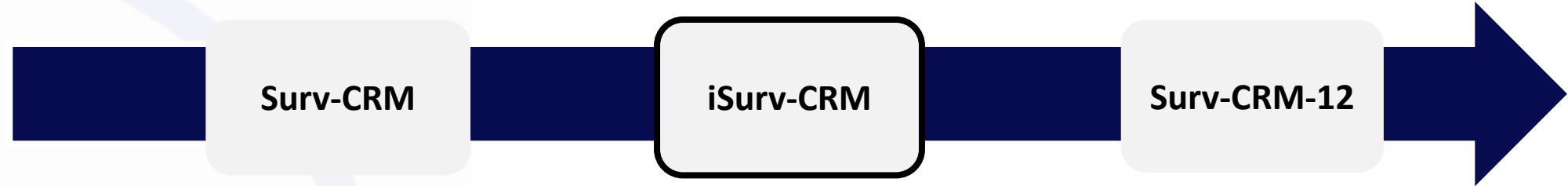
1. **Replacement** of non-evaluable patients
 2. Practical strategies for managing partial observation
 - **Weighted likelihood**
- Impacting performance of dose-toxicity centered designs (Biard et al 2020)



SURVIVAL-CONTINUAL REASSESSMENT METHOD (SURV-CRM)¹

- Handling fast and continual patient accrual
- Phase I design
 - Objective: identify the **MTD**

¹ Andrillon et al. *JBS* 2020



INFORMATIVE SURVIVAL-CONTINUAL REASSESSMENT METHOD (SURV-CRM)¹

- Handling fast and continual patient accrual
- **Early treatment discontinuation**

Phase I design

- Objective: identify the **MTD**

¹ Andrillon et al. *JBS* 2020



SURVIVAL-CONTINUAL REASSESSMENT METHOD-12 (SURV-CRM-12)²

Disease progression

- Reported in approximately 70% of the cases of premature discontinuation from a phase I trial (Olmos et al. 2012)
- **Efficacy information, but requiring treatment discontinuation**

Phase I/II design

- Objective: identify the **Optimal Dose (OD)** based on toxicity and efficacy-related information

² Andrillon et al. *Stat Med* 2022

COMBOLA TRIAL*

- Patients with low-risk myelodysplastic syndrome (MDS)
- Dose finding trial of luspatercept in combination with erythropoietin
- Toxicity observation windows: 42 days
- 5 dose levels

- Phase I/II trial using the **TITE-BOIN-ET design** (Takeda et al. 2020)

- **Progression free survival (PFS) was a secondary endpoint**

*Groupe Francophone des Myélodysplasies (GFM), NCT05181735

- **Toxicity and progression endpoints** within the observation window (t^*)
 - **T , the time-to-DLT or progression**
- **Administrative censoring**: during the trial and at the end of the observation window
- **Informative censoring**: discontinuations due to progression

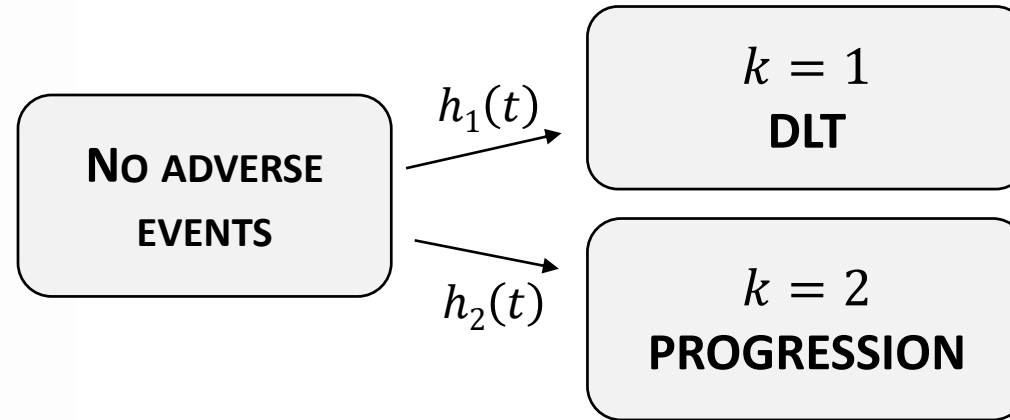
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- Survival analysis framework with competing risks
- $h_1(\cdot)$, the **cause-specific instantaneous hazard function** for DLT, d scaled dose
- $h_2(\cdot)$, the **cause-specific instantaneous hazard function**, for progression, d scaled dose

$$h_1(\beta_1, d) = \exp(d \exp(\beta_1))$$

$$h_2(\beta_2, d) = \exp(-d \exp(\beta_2))$$

*Benichou & Gail. *Biometrics* 1990



Observed cumulative incidence for event $k = 1, 2$ at time t^*

$$F_k(t^*, h_1, h_2, d) = \frac{h_k(t^*, d)}{h_1(t^*, d) + h_2(t^*, d)} \left(1 - \exp \left(- (H_1(t^*, d) + H_2(t^*, d)) \right) \right)$$

with $h_k(\cdot)$ the cause-specific hazard function,
and $H_k(\cdot)$ the corresponding cumulative hazards
event $k = 1$ for DLT and $k = 2$ for progression
 $[0; t^*]$: observation window

Dose finding objective

The optimal dose : $d^* = \underset{d \in \mathcal{A}_{\text{ccept}}}{\text{arg min}} F_2(t^*, h_1, h_2, d)$

$F_1(\cdot)$ the cumulative incidence of toxicity and $F_2(\cdot)$ the cumulative incidence of progression

Dose finding objective

The optimal dose : $d^* = \underset{d \in \mathcal{A}_{\text{accept}}}{\text{arg min}} F_2(t^*, h_1, h_2, d)$

$F_1(\cdot)$ the cumulative incidence of toxicity and $F_2(\cdot)$ the cumulative incidence of progression

Dose finding algorithm

$\mathcal{A}_{\text{accept}}$, the set of *acceptable doses*, $\mathcal{A} = \{d : d \leq \underset{d \in D}{\text{arg min}} |F_1(t^*, h_1, h_2, d) - \pi_{DLT}|\}$

$\mathcal{G}_{\text{good}}$, the set of *good doses*, $\mathcal{G} = \{d \in \mathcal{A} : F_2(t^*, h_1, h_2, d) - \min (F_2(t^*, h_1, h_2, d)) \leq \delta_p, \delta_p \geq 0\}$

Adaptive randomization

Randomization of the next inclusion at dose $d \in \mathcal{G}$ with probability $R = \frac{1 - F_2(t^*, h_1, h_2, d)}{\sum_{d \in \mathcal{G}_{\text{good}}} 1 - F_2(t^*, h_1, h_2, d)}$

DATA

- C : Right-censored failure time
- Y : Observed outcome
 - $Y=1$ if DLT, $Y=2$ if Progression, $Y=0$ otherwise
- D : Allocated doses

PRIOR DISTRIBUTIONS OF β_k , $\phi(\beta_k)$

Normal distribution with mean 0 and least informative prior variance*



Inference on cause-specific hazard performed separately for toxicity and progression**

Posterior mean of β_k : $\hat{\beta}_k = \frac{\int_{-\infty}^{\infty} \beta_k L(B;C,Y,D) \phi(\beta_k) d\beta_k}{\int_{-\infty}^{\infty} L(B;C,Y,D) \phi(\beta_k) d\beta_k}$, with $L(B;C, Y, D)$, the survival likelihood



Estimates of the cumulative incidences of DLT and Progression: $\hat{F}_k(t^*, h_1, h_2, d)$ for each dose level



Apply the dose-finding objective

* Lee and Cheung. *Stat in Med* 2011; **Benichou & Gail. *Biometrics* 1990

- **TITE-BOIN-ET** (Takeda et al. 2020)
 - Handling pending data via a weighted likelihood
 - Obj: the dose maximizing the efficacy probability among the doses lower or equal to the MTD

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Incomplete information (*Real clinical trial*) {

	Doses				
	D1	D2	D3	D4	D5
Y_1	*	*	0	*	*

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		Doses					
		D1	D2	D3	D4	D5	
Incomplete information (<i>Real clinical trial</i>)	{	Y_1	0	0	0	*	*
Complete information (<i>Framework of simulations</i>)	{	Y_1	0	0	0	1	1

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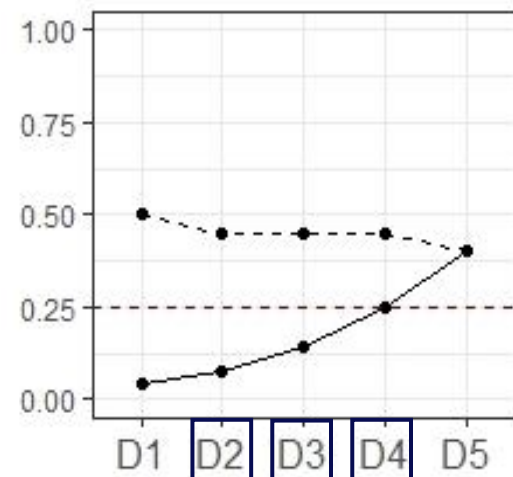
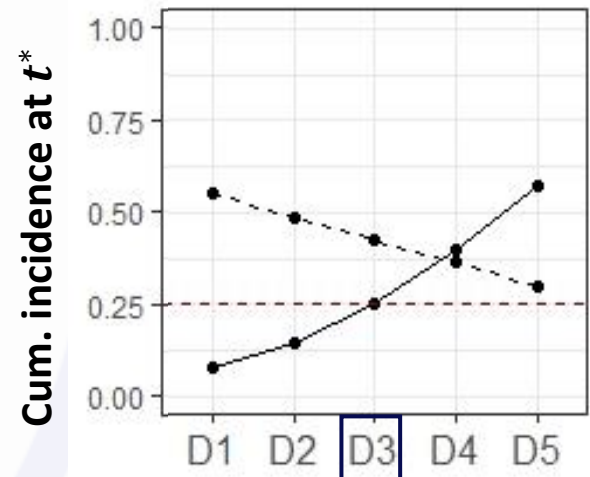
→ Provide a scenario-specific assessment of the accuracy of dose-finding designs
(PCS upper bound estimate)

SIMULATION STUDY: SCENARIOS

N = 10,000 trials; n = 45 patients; $\pi_{DLT} = 0.25$
Accrual rate of 4 patients per obs. window

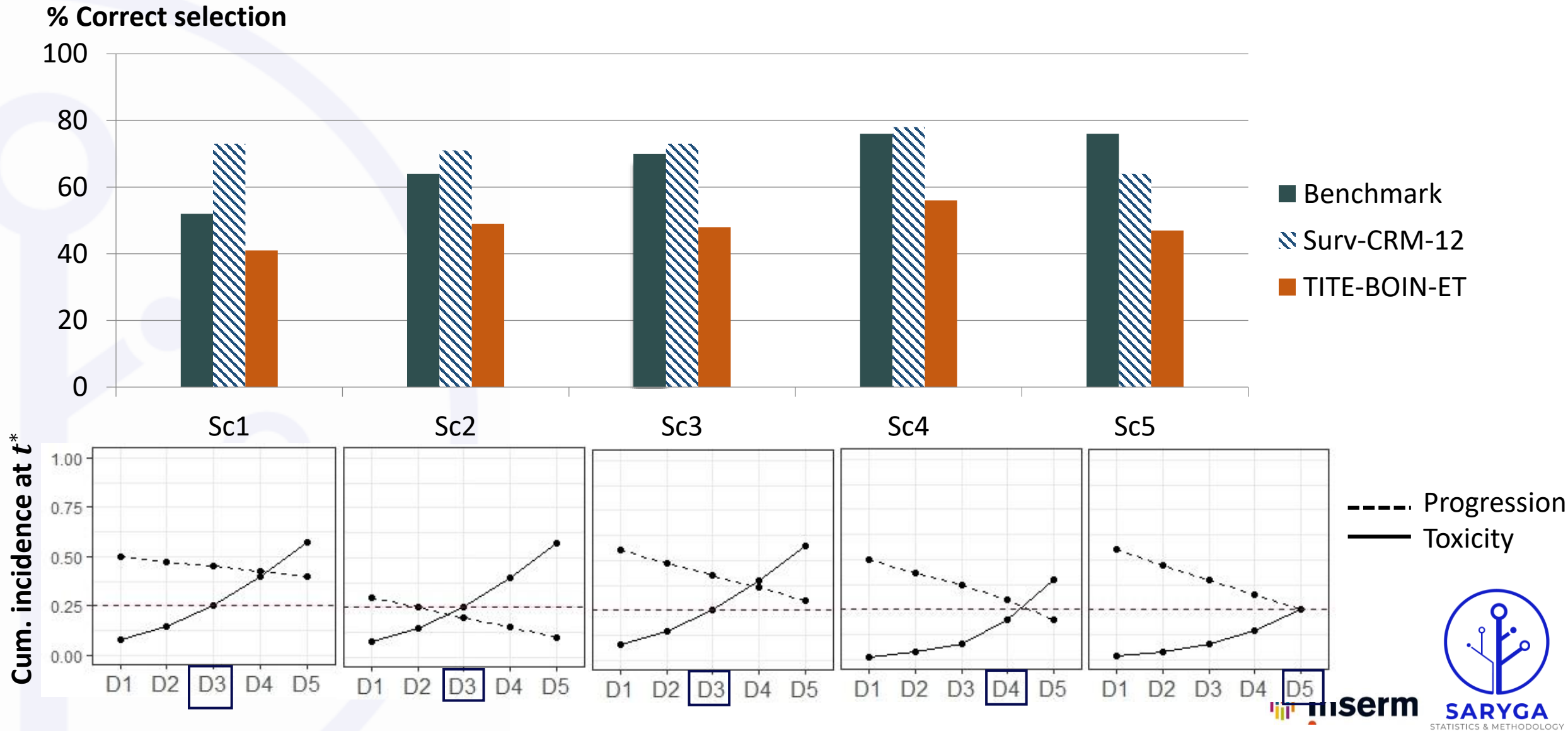
12 scenarios

- Cytotoxic
- Plateau dose progression relationship

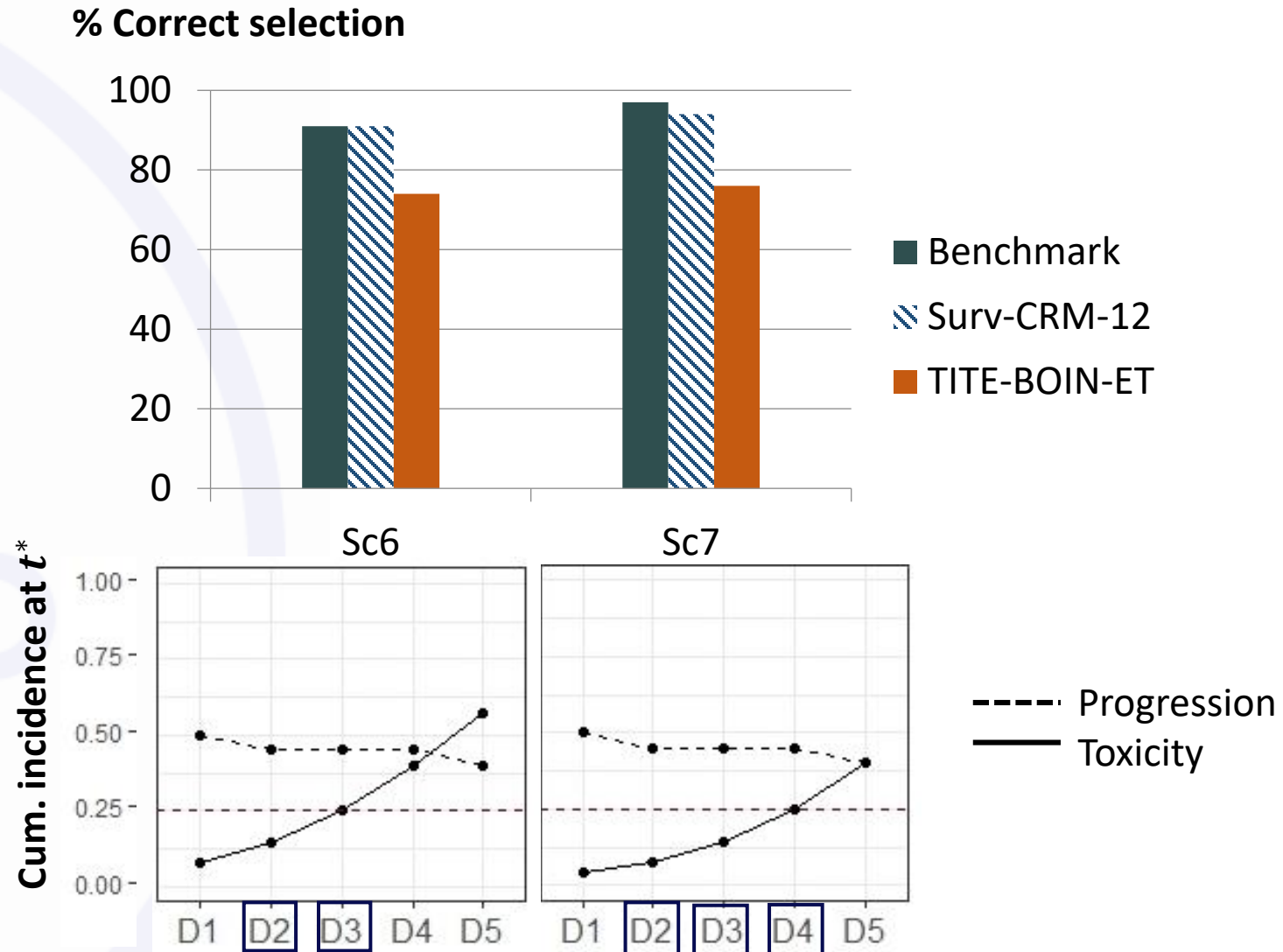


----- Progression
———— Toxicity

RESULTS: SURV-CRM-12; CYTOTOXIC SCENARIOS

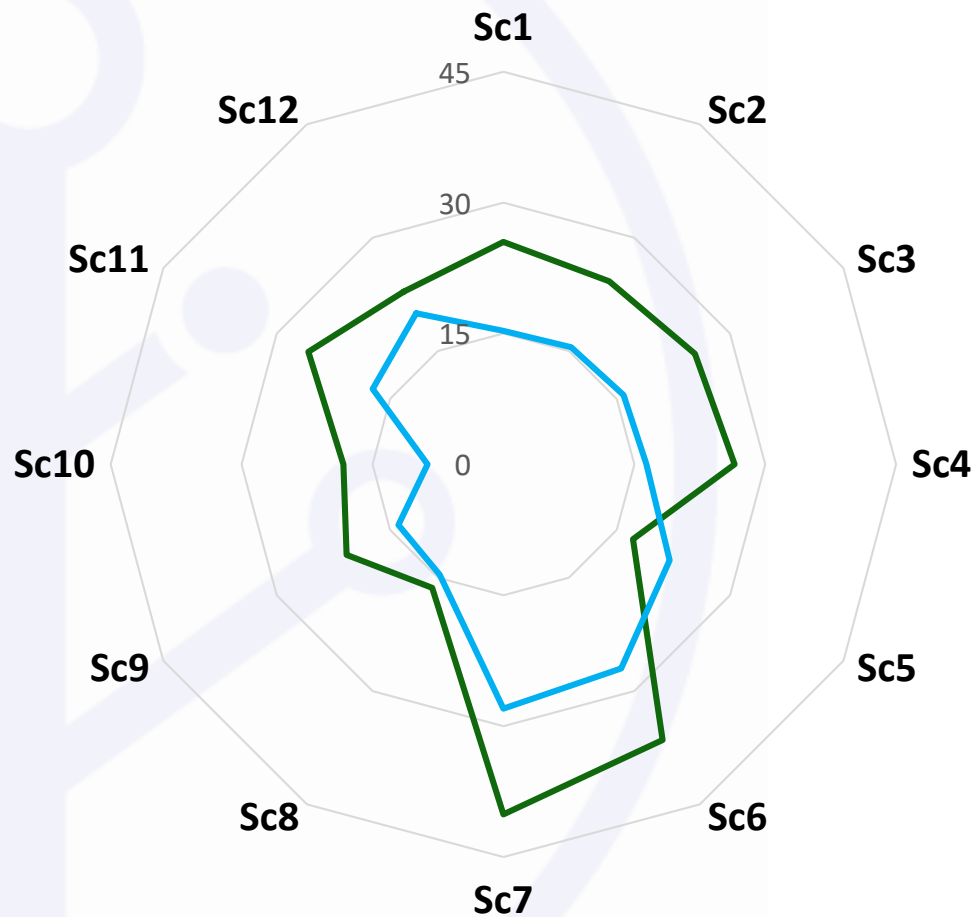


RESULTS: SURV-CRM-12; PLATEAU SCENARIOS



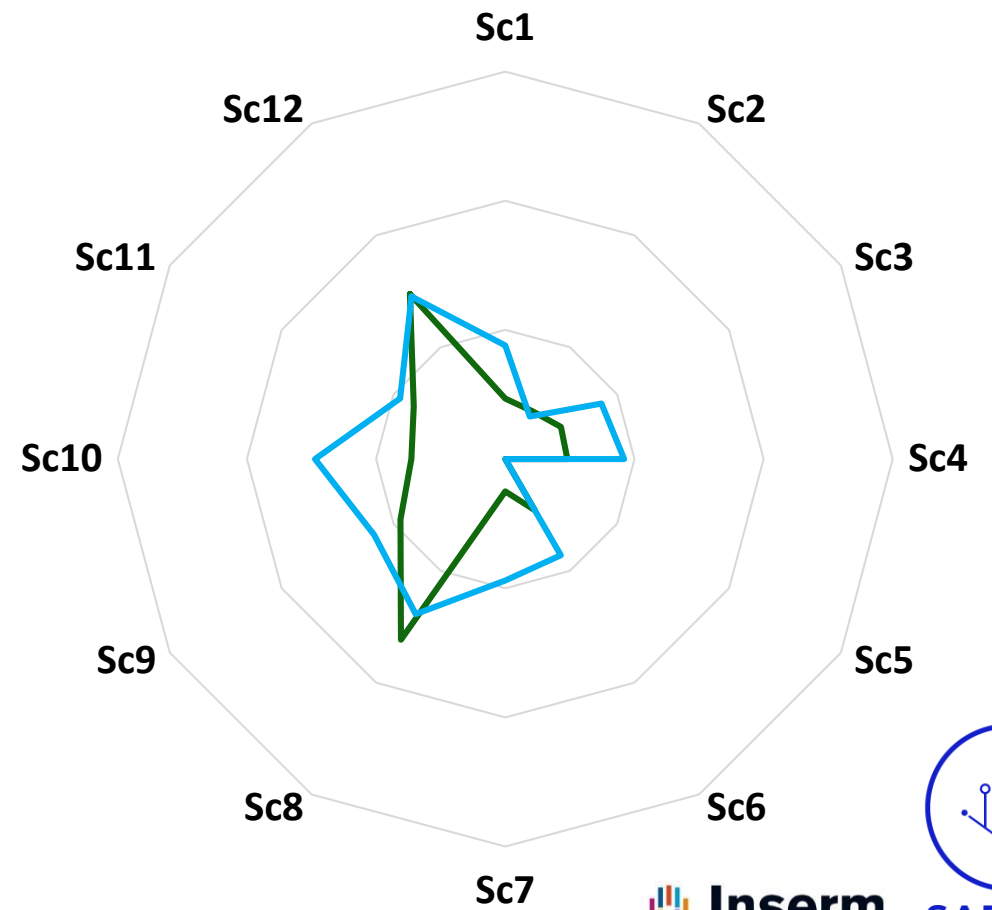
RESULTS: SURV-CRM-12; SAFETY DURING THE TRIAL

No. treated at the true OD



— Surv-CRM-12
— TITE-BOIN-ET

No. over treated



A FRAMEWORK FOR DOSE-FINDING (PHASE I & PHASE I/II)

WITH RIGHT CENSORED ENDPOINTS AND COMPETING RISKS ISSUES

→ Surv-CRM-12

- **Desirable properties** (statistical performance, safety, feasibility)
- **Sensitivity analyses** (patient accrual rate, sample size, variance specification, correlated time to toxicity and progression)



SARYGA
STATISTICS & METHODOLOGY

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- Jeong, J.-H. and Fine, J. (2006). Direct parametric inference for the cumulative incidence function. *Journal of the Royal Statistical Society: Series C (Applied Statistics)*; 55(2):187–200.
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- O’Quigley, J., Paoletti, X., and Maccario, J. (2002). Non-parametric optimal design in dose finding studies. *Biostatistics*; 3(1):51–56.
- Putter H., Fiocco M., Geskus R.B. (2006). Tutorial in biostatistics: competing risks and multi-state models. *Statistics in Medicine*; 26(11):2389-430.
- Zhang Y., S. Cao, et al. (2021). A Bayesian adaptive phase I/II clinical trial design with late-onset competing risk outcomes". *Biometrics*; 77 (3), pp. 796–808.

BACK-UP



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STATISTICS & METHODOLOGY

Marginal cause-specific cumulative incidences of event $k = 1, 2$ at time t^* (unobserved)

$$1 - \exp(H_k(t^*, d))$$

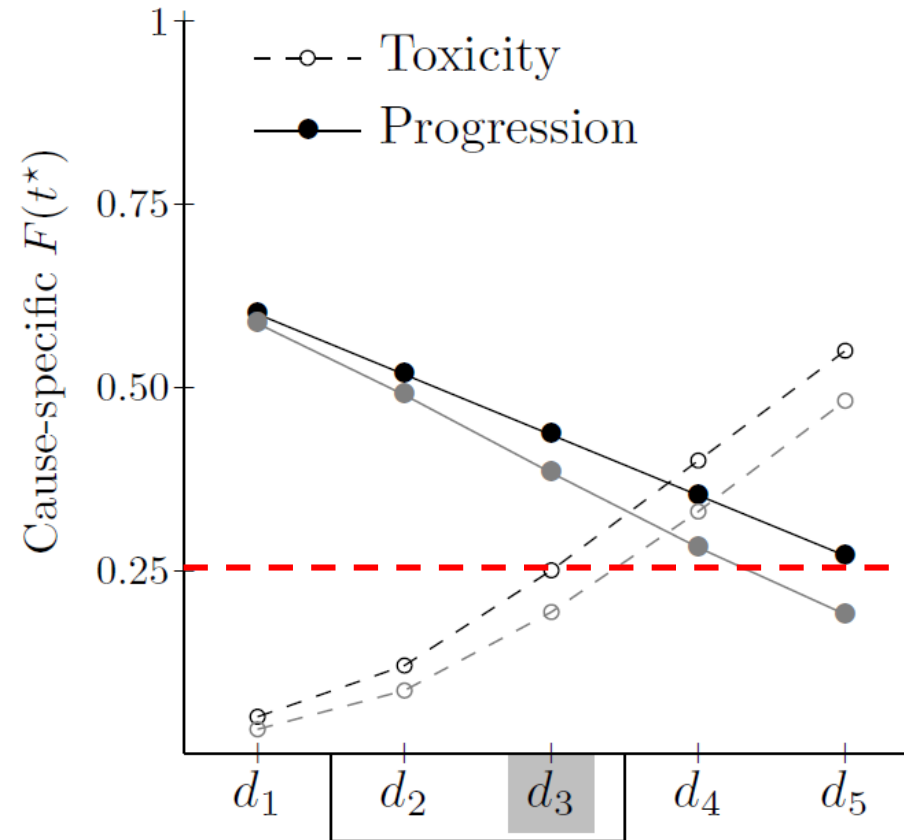
Observed cumulative incidence for event $k = 1, 2$ at time t^*

$$F_k(t^*, h_1, h_2, d) = \frac{h_k(t^*, d)}{h_1(t^*, d) + h_2(t^*, d)} \left(1 - \exp \left(- (H_1(t^*, d) + H_2(t^*, d)) \right) \right)$$

with $h_k(\cdot)$ the cause-specific hazard function,
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event $k = 1$ for DLT and $k = 2$ for treatment discontinuation
[0; t^*]: observation window

OBJECTIVE

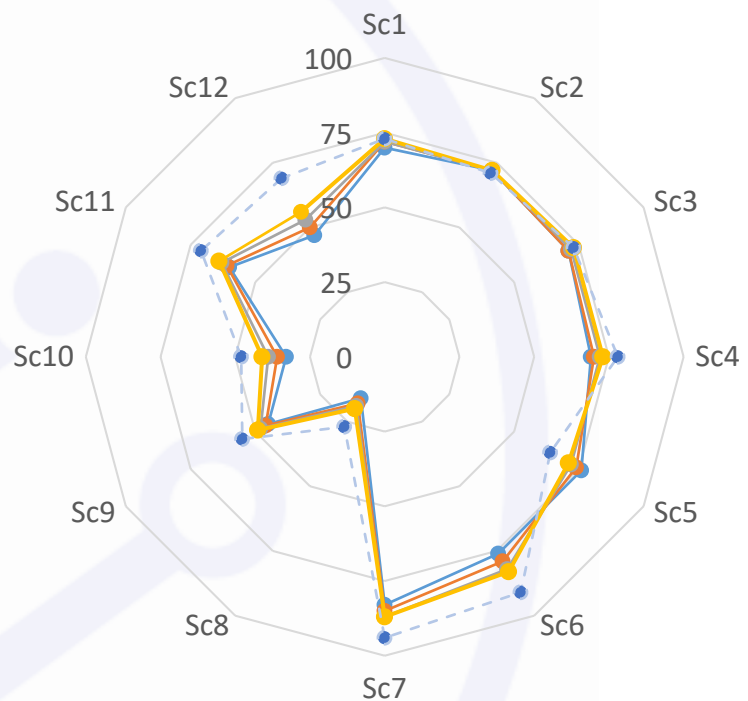
- Black lines (solid and dashed) : Cause-specific marginal cumulative incidences of event at t^*
- Gray lines (solid and dashed): Observed cumulative incidences of events at t^*



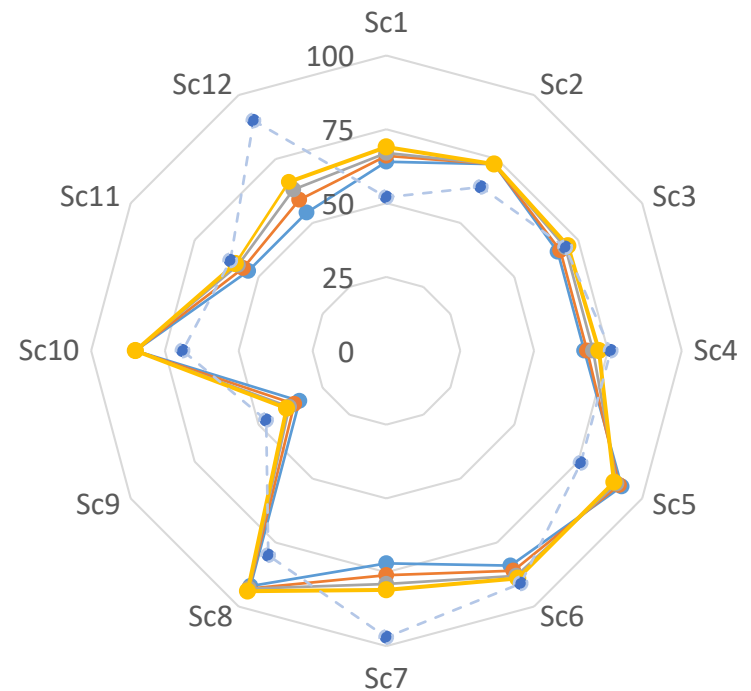
- **Competing risks data generation** (Beyersmann et al. 2009)
 - Cause-specific risks back computed from scenarios cumulative incidence at t^*
 - Time to any event sampled from exponential distribution with rate (h_1+h_2)
 - Event case determined by a random draw from a Bernoulli (h_1/h_1+h_2) for toxicity
 - Administrative censoring applied at t^*
- **Dose skeletons calibration:** indifference intervals (Lee and Cheung 2009)
- **Least informative prior variance for β_1 and β_2** (Lee and Cheung 2011)

RESULTS: SURV-CRM-12; SENSITIVITY ANALYSES

% Correct selection Surv-CRM-12



% Correct selection Benchmark



- c = 1
- c = 1.2
- c = 1.5
- c = 1.8
- Not correlated

Clayton model for data generation

When $c \rightarrow 0$ the correlation approaches 1 and, when $c \rightarrow \infty$ the correlation converges to 0

Competing Risk-CRM (Biard et al. 2021)

- Survival framework, targeting **theoretical marginal incidence of events**

Obj: a dose or a set of doses associated with the **minimum progression marginal risk within an admissible set**

Bayesian data augmentation design (Zhang et al. 2021)

- Late-onset competing risks outcome modeled by the **cause-specific hazard rate**
- Piecewise exponential model
- Pending data during the trial are treated as **missing data**

Obj: the dose yielding the **highest posterior mean utility within an admissible set**

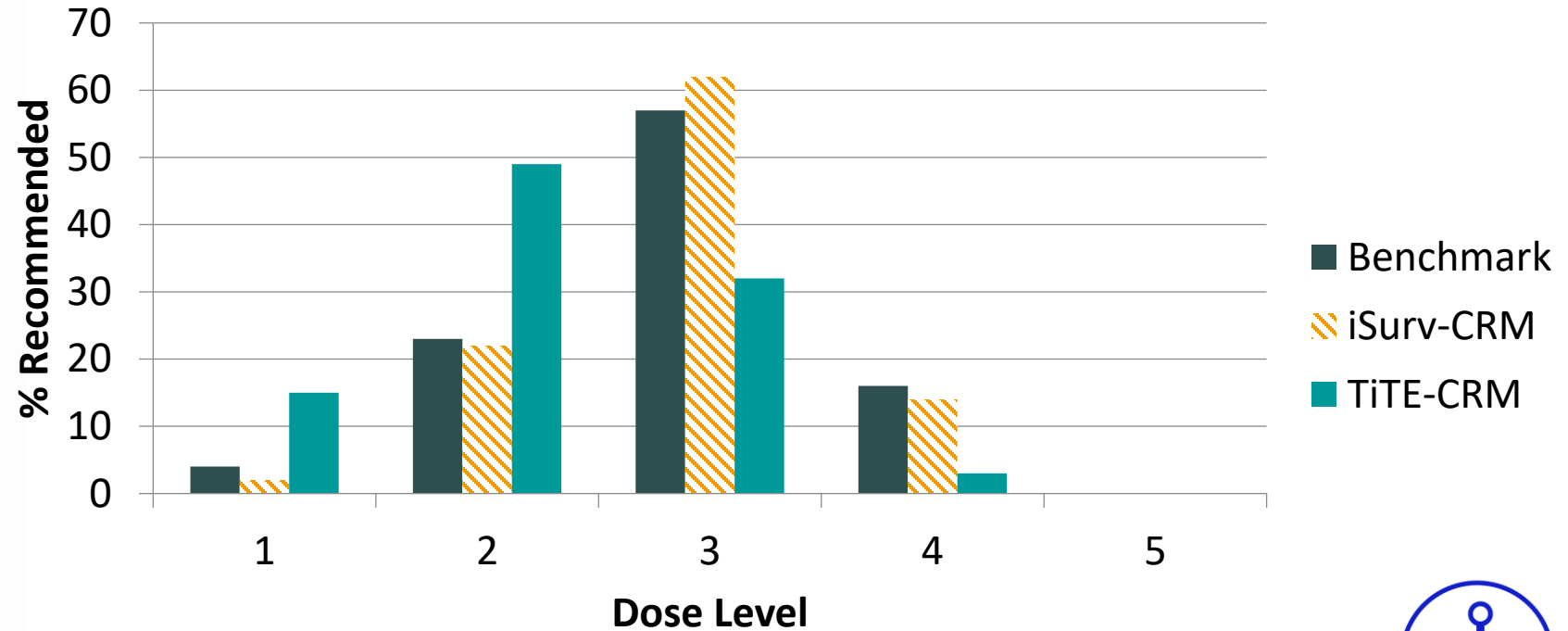
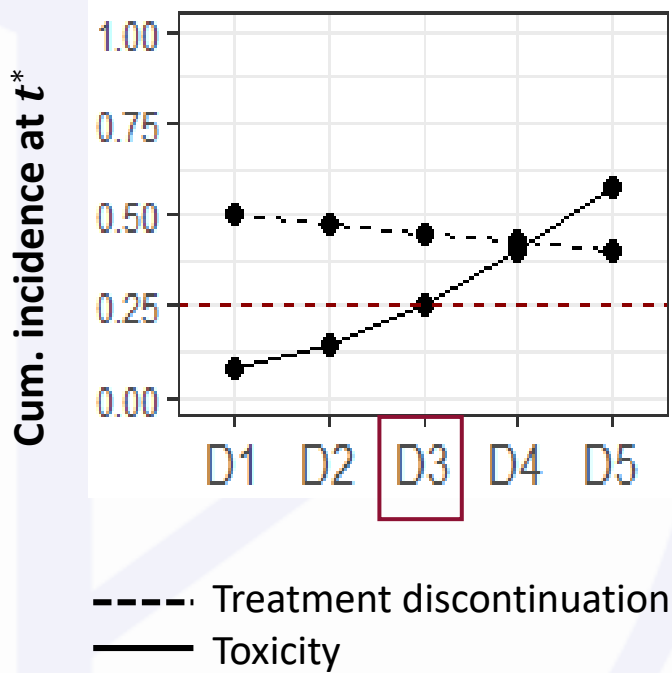
- **CR-CRM**
 - Likelihood estimation
 - 3 stage design
 - Target unobserved cause specific incidences
 - Non monotone dose progression relationship
- **Surv-CRM-12:**
 - Bayesian
 - One-stage design
 - Target observed cumulative incidence of events
 - Monotonicity assumption

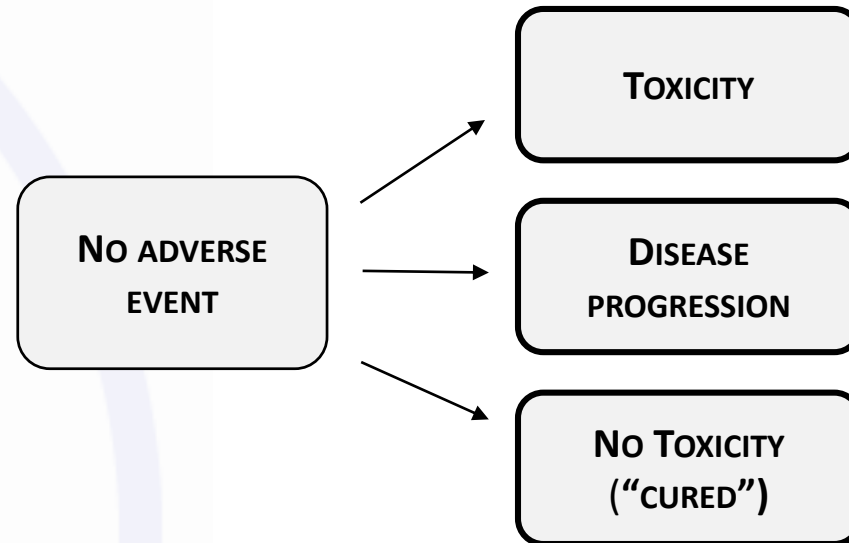
- Surv-CRM {0.069, 0.151, 0.250, 0.346, 0.426}.
- iSurv-CRM, best guess prior. Polley (2011),
 - Toxicity: {0.05, 0.10, 0.25, 0.35, 0.50}
 - Trial discontinuation: {0.50, 0.45, 0.40, 0.35, 0.35}
- TITE-CRM: {0.043; 0.124; 0.250; 0.398; 0.542}. Lee and Cheung (2009)
- **Surv-CRM-12**
 - Toxicity: {0.055, **0.130**, **0.250**, **0.406**, 0.571}
 - Progression {0.666, 0.541, 0.400, 0.266, 0.158}.

RESULTS: ISURV-CRM

N = 10,000 trials; n = 25 patients; $\pi_{DLT} = 0.25$
Accrual rate of 4 patients per obs. window

Scenario 1





- Patients who experienced neither disease progression nor DLT at t^* → *informative censoring*
 - **Trinomial response outcomes** (Zhang et al. 2021)
 - **Cure rate models:** a fraction of subjects in the population will never experience the event of interest (Berkson and Gage, 2006)

HANDLING ADMINISTRATIVE CENSORING

1. Weight the observations

	Model-based	Model-assisted
Phase I	TITE-CRM (Cheung and Chappell 2000)	TITE-mTPI (Lin and Yuan 2020)
Phase I/II	Seamless phase I/II TITE-CRM (Yan et al. 2019), Robust Bayesian EffTox design (Liu and Johnson 2016) Phase I/II MTA (Riviere et al. 2016)	TITE-BOIN-ET (Takeda et al. 2020)

2. Time-to-event endpoint

Phase I	fCRM (Yin et al. 2013)	
Phase I/II	Survival joint model (Yuan and Yin, 2009), Survival EffTox (Koopmeiners and Modiano 2014)	

3. Missing data methodology

Phase I	EM-CRM (Yuan and Yin 2011), DA-CRM (Liu et al. 2013)	TITE-BOIN (Yuan et al. 2018)
Phase I/II	LO-EffTox (Jin et al. 2014)	TITE-BOIN-12 (Zhou et al. 2022)

- **Toxicity endpoint:** DLT within the observation window (t^*)
 - **T , the time-to-DLT**
- **Administrative censoring:** during the trial and at the end of the observation window
- Independent from the time-to-DLT

- Survival analysis framework
 - **Instantaneous hazard of toxicity** for DLT, d scaled dose: $h(\beta, d) = \exp(d \exp(\beta))$
 - Cumulative hazard, $H(\beta, d)$
 - **Cumulative incidence of DLT** at the end of the observation window, t^*

$$F(d, t^*, h) = 1 - \exp(-h(\beta, d)t^*)$$

- CRM dose finding-algorithm

Identify the MTD, d^* , the dose with the probability of toxicity at time t^* closest to a pre-specified target π_{DLT} (e.g., 0.25), among the set D of candidate dose levels

$$d^* = \mathit{arg\,min}_{d \in D} |F(d, t^*, h) - \pi_{DLT}|$$

- With $F(d, t^*, h)$, the **cumulative incidence of toxicity** at time t^*

- CRM dose finding-algorithm

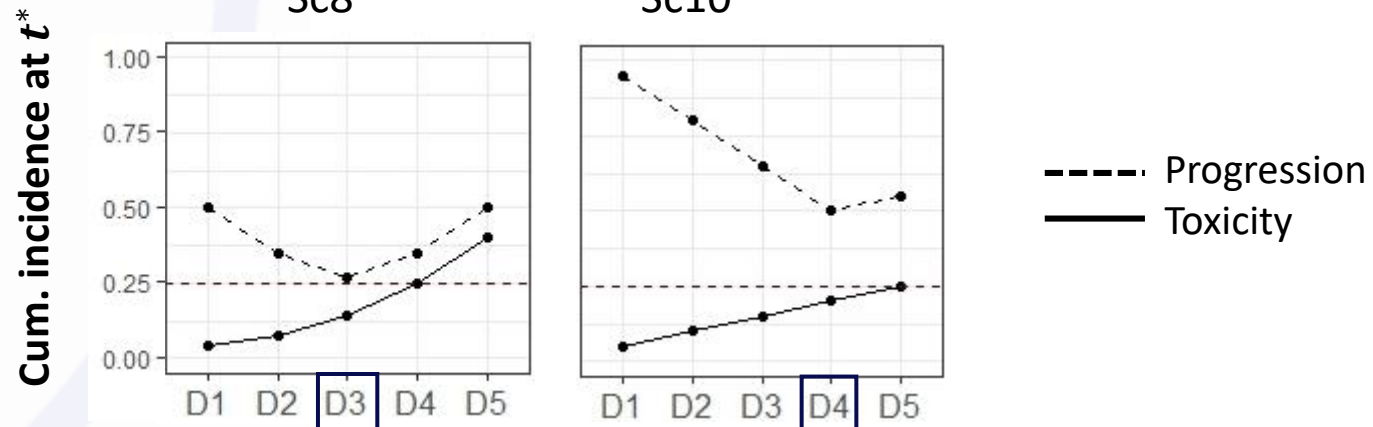
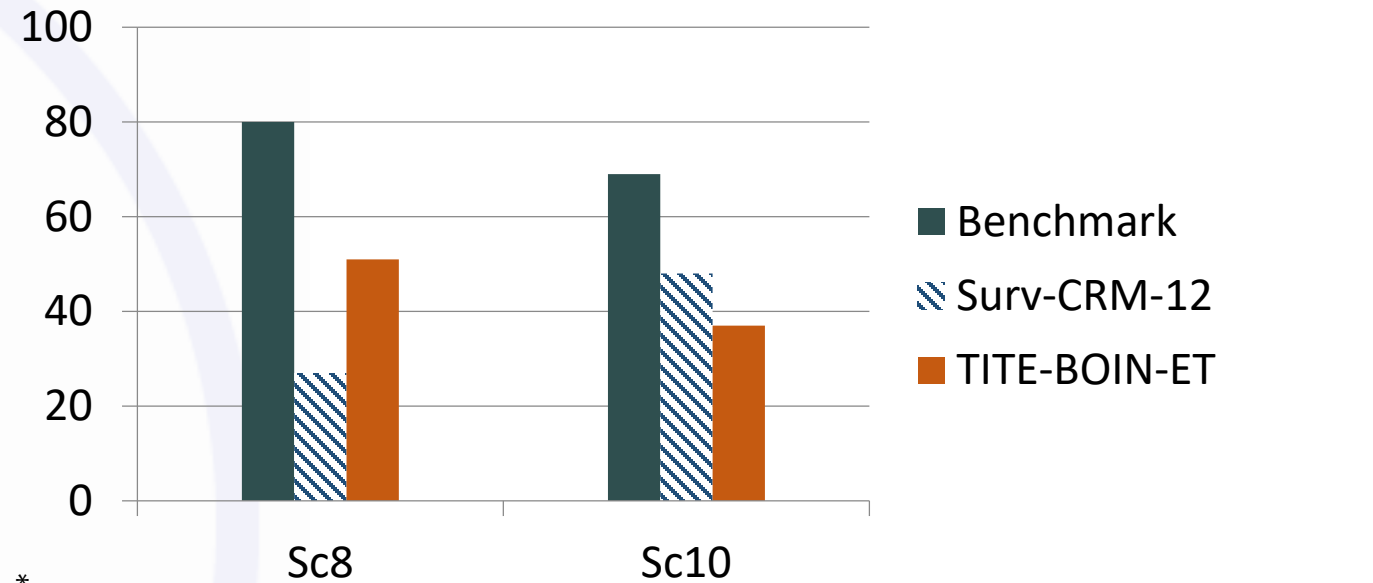
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$$d^* = \mathit{arg\,min}_{d \in D} |F_1(t^*, h_1, h_2, d) - \pi_{DLT}|$$

- With $F_1(\cdot)$, the **observed cumulative incidence for toxicity** at time t^*

RESULTS: SURV-CRM-12; U-SHAPE SCENARIOS

% Correct selection



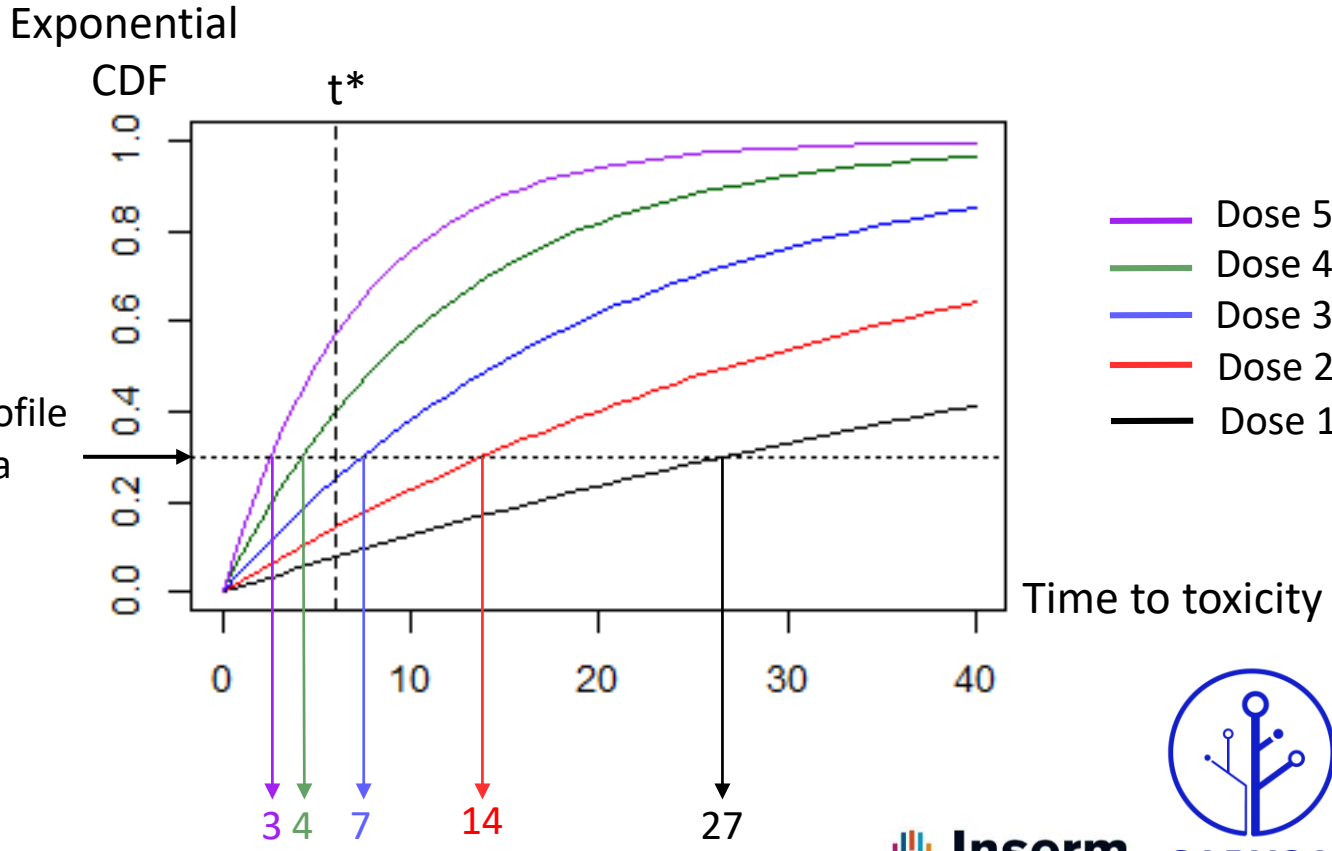
SIMULATION STUDY: COMPARATORS

Nonparametric Benchmark (O'Quigley et al. 2002, Cheung 2014, Mozgunov et al. 2020)

specific assessment of the accuracy of
 PCS upper bound estimate

	Doses					
		D1	D2	D3	D4	D5
<i>(Real clinical trial)</i>	Y_1	0	0	0	*	*
<i>(Simulations)</i>	Y_1	0	0	0	1	1

Tolerance profile drawn from a Uniform(1,0)



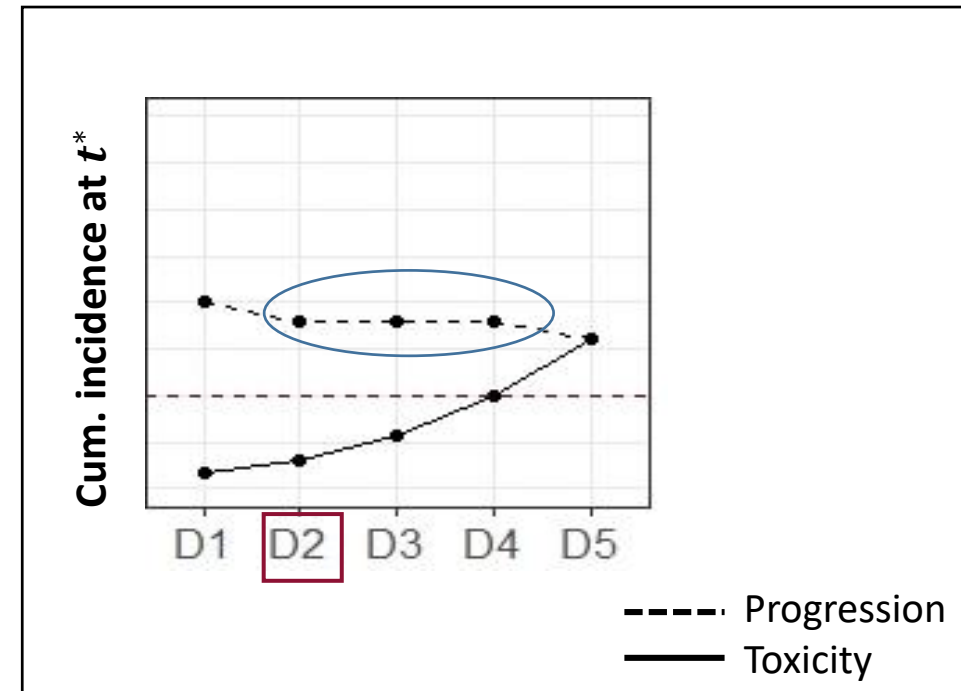
True Toxicity probability by dose : {0.08, 0.14, 0.25, 0.40, 0.57}

Surv-CRM-12 Optimal Dose definition: $d^* = \arg \min_{d \in \mathcal{A}} F_2(t^*, h_1, h_2, d)$

- 'Optimus' FDA project recommendations
 - **OD: lowest** safe dose that achieves the highest efficacy
- **Extension of the efficacy working model to strictly non-monotone relationship**

Cause-specific instantaneous hazard function, for progression

$$h_2(\beta_2, d) = \exp(\beta_{02} + \exp(\beta_{12}) d + \exp(\beta_{22}) d^2)$$



Intention To Treat (ITT) approach

- Treatment decision at the time of the patient's inclusion in the trial

- **Repeated treatment administration over several cycles**

- Estimation of the MTD associated to some cumulative risk of DLT over a predefined number of treatment cycles (Altzerinakou, 2019, Ursino, 2021)